

# Chapter 19

## The SYSLIN Procedure

### Chapter Table of Contents

---

<b>OVERVIEW</b> . . . . .	1051
<b>GETTING STARTED</b> . . . . .	1052
An Example Model . . . . .	1052
Variables in a System of Equations . . . . .	1053
Using PROC SYSLIN . . . . .	1053
OLS Estimation . . . . .	1054
Two-Stage Least Squares Estimation . . . . .	1055
LIML, K-Class, and MELO Estimation . . . . .	1057
SUR, 3SLS, and FIML Estimation . . . . .	1058
Computing Reduced Form Estimates . . . . .	1061
Restricting Parameter Estimates . . . . .	1063
Testing Parameters . . . . .	1065
Saving Residuals and Predicted Values . . . . .	1068
Plotting Residuals . . . . .	1068
<b>SYNTAX</b> . . . . .	1070
Functional Summary . . . . .	1070
PROC SYSLIN Statement . . . . .	1072
BY Statement . . . . .	1075
ENDOGENOUS Statement . . . . .	1075
IDENTITY Statement . . . . .	1075
INSTRUMENTS Statement . . . . .	1075
MODEL Statement . . . . .	1076
OUTPUT Statement . . . . .	1077
RESTRICT Statement . . . . .	1078
SRESTRICT Statement . . . . .	1079
STEST Statement . . . . .	1080
TEST Statement . . . . .	1082
VAR Statement . . . . .	1083
WEIGHT Statement . . . . .	1083
<b>DETAILS</b> . . . . .	1084
Input Data Set . . . . .	1084
Estimation Methods . . . . .	1084
ANOVA Table for Instrumental Variables Methods . . . . .	1087

*Part 2. General Information*

The $R^2$ Statistics . . . . .	1087
Computational Details . . . . .	1088
Missing Values . . . . .	1090
OUT= Data Set . . . . .	1090
OUTEST= Data Set . . . . .	1090
OUTSSCP= Data Set . . . . .	1092
Printed Output . . . . .	1092
ODS Table Names . . . . .	1094
<b>EXAMPLES</b> . . . . .	1096
Example 19.1 Klein's Model I Estimated with LIML and 3SLS . . . . .	1096
Example 19.2 Grunfeld's Model Estimated with SUR . . . . .	1103
<b>REFERENCES</b> . . . . .	1108

# Chapter 19

## The SYSLIN Procedure

---

### Overview

The SYSLIN procedure estimates parameters in an interdependent system of linear regression equations.

Ordinary least squares (OLS) estimates are biased and inconsistent when current period endogenous variables appear as regressors in other equations in the system. The errors of a set of related regression equations are often correlated, and the efficiency of the estimates can be improved by taking these correlations into account. The SYSLIN procedure provides several techniques which produce consistent and asymptotically efficient estimates for systems of regression equations.

The SYSLIN procedure provides the following estimation methods:

- ordinary least squares (OLS)
- two-stage least squares (2SLS)
- limited information maximum likelihood (LIML)
- K-class
- seemingly unrelated regressions (SUR)
- iterated seemingly unrelated regressions (ITSUR)
- three-stage least squares (3SLS)
- iterated three-stage least squares (IT3SLS)
- full information maximum likelihood (FIML)
- minimum expected loss (MELO)

Other features of the SYSLIN procedure enable you to:

- impose linear restrictions on the parameter estimates.
- test linear hypotheses about the parameters.
- write predicted and residual values to an output SAS data set.
- write parameter estimates to an output SAS data set.
- write the crossproducts matrix (SSCP) to an output SAS data set.
- use raw data, correlations, covariances, or cross products as input.

---

## Getting Started

This section introduces the use of the SYSLIN procedure. The problem of dependent regressors is introduced using a supply-demand example. This section explains the terminology used for variables in a system of regression equations and introduces the SYSLIN procedure statements for declaring the roles the variables play. The syntax used for the different estimation methods and the output produced is shown.

---

### An Example Model

In simultaneous systems of equations, endogenous variables are determined jointly rather than sequentially. Consider the following demand and supply functions for some product:

$$Q_D = a_1 + b_1P + c_1Y + d_1S + \epsilon_1 \quad (\text{demand})$$

$$Q_S = a_2 + b_2P + c_2U + \epsilon_2 \quad (\text{supply})$$

$$Q = Q_D = Q_S \quad (\text{market equilibrium})$$

The variables in this system are as follows:

$Q_D$	quantity demanded
$Q_S$	quantity supplied
$Q$	the observed quantity sold, which equates quantity supplied and quantity demanded in equilibrium
$P$	price per unit
$Y$	income
$S$	price of substitutes
$U$	unit cost
$\epsilon_1$	the random error term for the demand equation
$\epsilon_2$	the random error term for the supply equation

In this system, quantity demanded depends on price, income, and the price of substitutes. Consumers normally purchase more of a product when prices are lower and when income and the price of substitute goods are higher. Quantity supplied depends on price and the unit cost of production. Producers will supply more when price is high and when unit cost is low. The actual price and quantity sold are determined jointly by the values that equate demand and supply.

Since price and quantity are jointly endogenous variables, both structural equations are necessary to adequately describe the observed values. A critical assumption of OLS is that the regressors are uncorrelated with the residual. When current endogenous variables appear as regressors in other equations (endogenous variables depend

on each other), this assumption is violated and the OLS parameter estimates are biased and inconsistent. The bias caused by the violated assumptions is called *Simultaneous equation bias*. Neither the demand nor supply equation can be estimated consistently by OLS.

---

## Variables in a System of Equations

Before explaining how to use the SYSLIN procedure, it is useful to define some terms. The variables in a system of equations can be classified as follows:

- *Endogenous variables*, which are also called *jointly dependent* or *response variables*, are the variables determined by the system. Endogenous variables can also appear on the right-hand side of equations.
- *Exogenous variables* are independent variables that do not depend on any of the endogenous variables in the system.
- *Predetermined variables* include both the exogenous variables and *lagged endogenous variables*, which are past values of endogenous variables determined at previous time periods. PROC SYSLIN does not compute lagged values; any lagged endogenous variables must be computed in a preceding DATA step.
- *Instrumental variables* are predetermined variables used in obtaining predicted values for the current period endogenous variables by a first-stage regression. The use of instrumental variables characterizes estimation methods such as two-stage least squares and three-stage least squares. Instrumental variables estimation methods substitute these first-stage predicted values for endogenous variables when they appear as regressors in model equations.

---

## Using PROC SYSLIN

First specify the input data set and estimation method on the PROC SYSLIN statement. If any model uses dependent regressors, and you are using an instrumental variables regression method, declare the dependent regressors with an ENDOGENOUS statement and declare the instruments with an INSTRUMENTS statement. Next, use MODEL statements to specify the structural equations of the system.

The use of different estimation methods is shown by the following examples. These examples use simulated data (not shown).

## OLS Estimation

PROC SYSLIN performs OLS regression if you do not specify a method of estimation in the PROC SYSLIN statement. OLS does not use instruments, so the ENDOGENOUS and INSTRUMENTS statements can be omitted.

The following statements estimate the supply and demand model shown previously:

```
proc syslin data=in;
  demand: model q = p y s;
  supply: model q = p u;
run;
```

The PROC SYSLIN output for the demand equation is shown in Figure 19.1, and the output for the supply equation is shown in Figure 19.2.

The SYSLIN Procedure						
Ordinary Least Squares Estimation						
Model	DEMAND					
Dependent Variable	q					
Label	Quantity					
Analysis of Variance						
Source	DF	Sum of Squares	Mean Square	F Value	Pr > F	
Model	3	9.587891	3.195964	398.31	<.0001	
Error	56	0.449336	0.008024			
Corrected Total	59	10.03723				
Root MSE		0.08958	R-Square	0.95523		
Dependent Mean		1.30095	Adj R-Sq	0.95283		
Coeff Var		6.88541				
Parameter Estimates						
Variable	DF	Parameter Estimate	Standard Error	t Value	Pr >  t	Variable Label
Intercept	1	-0.47677	0.210239	-2.27	0.0272	Intercept
p	1	0.123324	0.105177	1.17	0.2459	Price
y	1	0.201282	0.032403	6.21	<.0001	Income
s	1	0.167258	0.024091	6.94	<.0001	Price of Substitutes

Figure 19.1. OLS Results for Demand Equation

```

The SYSLIN Procedure
Ordinary Least Squares Estimation

Model
Dependent Variable   SUPPLY
Label                q
Quantity

Analysis of Variance

Source                DF      Sum of      Mean
                    Squares    Square    F Value    Pr > F

Model                 2      9.033890    4.516945    256.61    <.0001
Error                57      1.003337    0.017602
Corrected Total      59      10.03723

Root MSE              0.13267    R-Square    0.90004
Dependent Mean        1.30095    Adj R-Sq    0.89653
Coeff Var             10.19821

Parameter Estimates

Variable              Parameter Standard      Variable
                    DF Estimate  Error t Value Pr > |t| Label

Intercept            1  -0.30390  0.471397   -0.64   0.5217 Intercept
p                    1   1.218743  0.053914   22.61  <.0001 Price
u                    1  -1.07757  0.234150   -4.60  <.0001 Unit Cost

```

**Figure 19.2.** OLS Results for Supply Equation

For each MODEL statement, the output first shows the model label and dependent variable name and label. This is followed by an Analysis of Variance table for the model, which shows the model, error, and total mean squares, and an  $F$  test for the no-regression hypothesis. Next, the procedure prints the root mean square error, dependent variable mean and coefficient of variation, and the  $R^2$  and adjusted  $R^2$  statistics.

Finally, the table of parameter estimates shows the estimated regression coefficients, standard errors, and  $t$ -tests. You would expect the price coefficient in a demand equation to be negative. However, note that the OLS estimate of the price coefficient P in the demand equation (.1233) has a positive sign. This could be caused by simultaneous equation bias.

---

## Two-Stage Least Squares Estimation

In the supply and demand model, P is an endogenous variable, and consequently the OLS estimates are biased. The following example estimates this model using two-stage least squares.

Part 2. General Information

```
proc syslin data=in 2sls;
  endogenous p;
  instruments y u s;
  demand: model q = p y s;
  supply: model q = p u;
run;
```

The 2SLS option on the PROC SYSLIN statement specifies the two-stage least-squares method. The ENDOGENOUS statement specifies that P is an endogenous regressor for which first-stage predicted values are substituted. You only need to declare an endogenous variable in the ENDOGENOUS statement if it is used as a regressor; thus although Q is endogenous in this model, it is not necessary to list it in the ENDOGENOUS statement.

Usually, all predetermined variables that appear in the system are used as instruments. The INSTRUMENTS statement specifies that the exogenous variables Y, U, and S are used as instruments for the first-stage regression to predict P.

The 2SLS results are shown in Figure 19.3 and Figure 19.4. The first-stage regressions are not shown. To see the first-stage regression results, use the FIRST option on the MODEL statement.

The SYSLIN Procedure						
Two-Stage Least Squares Estimation						
Model	DEMAND					
Dependent Variable	q					
Label	Quantity					
Analysis of Variance						
Source	DF	Sum of Squares	Mean Square	F Value	Pr > F	
Model	3	9.670882	3.223627	115.58	<.0001	
Error	56	1.561944	0.027892			
Corrected Total	59	10.03723				
Root MSE		0.16701	R-Square	0.86095		
Dependent Mean		1.30095	Adj R-Sq	0.85350		
Coeff Var		12.83740				
Parameter Estimates						
Variable	DF	Parameter Estimate	Standard Error	t Value	Pr >  t	Variable Label
Intercept	1	1.901040	1.171224	1.62	0.1102	Intercept
p	1	-1.11518	0.607391	-1.84	0.0717	Price
y	1	0.419544	0.117954	3.56	0.0008	Income
s	1	0.331475	0.088472	3.75	0.0004	Price of Substitutes

Figure 19.3. 2SLS Results for Demand Equation



The SYSLIN Procedure						
Two-Stage Least Squares Estimation						
Model			SUPPLY			
Dependent Variable			q			
Label			Quantity			
Analysis of Variance						
Source	DF	Sum of Squares	Mean Square	F Value	Pr > F	
Model	2	9.646098	4.823049	253.96	<.0001	
Error	57	1.082503	0.018991			
Corrected Total	59	10.03723				
Root MSE		0.13781	R-Square	0.89910		
Dependent Mean		1.30095	Adj R-Sq	0.89556		
Coeff Var		10.59291				
Parameter Estimates						
Variable	DF	Parameter Estimate	Standard Error	t Value	Pr >  t	Variable Label
Intercept	1	-0.51878	0.490999	-1.06	0.2952	Intercept
p	1	1.333080	0.059271	22.49	<.0001	Price
u	1	-1.14623	0.243491	-4.71	<.0001	Unit Cost

**Figure 19.4.** 2SLS Results for Supply Equation

The 2SLS output is similar in form to the OLS output. However, the 2SLS results are based on predicted values for the endogenous regressors from the first stage instrumental regressions. This makes the analysis of variance table and the  $R^2$  statistics difficult to interpret. See the sections "ANOVA Table for Instrumental Variables Methods" and "The  $R^2$  Statistics" later in this chapter for details.

Note that, unlike the OLS results, the 2SLS estimate for the P coefficient in the demand equation (-1.115) is negative.

## LIML, K-Class, and MELO Estimation

To obtain limited information maximum likelihood, general K-class, or minimum expected loss estimates, use the ENDOGENOUS, INSTRUMENTS, and MODEL statements as in the 2SLS case but specify the LIML, K=, or MELO option instead of 2SLS in the PROC SYSLIN statement. The following statements show this for K-class estimation.

```
proc syslin data=in k=.5;
  endogenous p;
  instruments y u s;
  demand: model q = p y s;
  supply: model q = p u;
run;
```

For more information on these estimation methods see the "Estimation Methods" in the "Details" section and consult econometrics textbooks.

---

## SUR, 3SLS, and FIML Estimation

In a multivariate regression model, the errors in different equations may be correlated. In this case the efficiency of the estimation may be improved by taking these cross-equation correlations into account.

### **Seemingly Unrelated Regression**

Seemingly unrelated regression (SUR), also called joint generalized least squares (JGLS) or Zellner estimation, is a generalization of OLS for multi-equation systems. Like OLS, the SUR method assumes that all the regressors are independent variables, but SUR uses the correlations among the errors in different equations to improve the regression estimates. The SUR method requires an initial OLS regression to compute residuals. The OLS residuals are used to estimate the cross-equation covariance matrix.

The SUR option on the PROC SYSLIN statement specifies seemingly unrelated regression, as shown in the following statements:

```
proc syslin data=in sur;  
  demand: model q = p y s;  
  supply: model q = p u;  
run;
```

INSTRUMENTS and ENDOGENOUS statements are not needed for SUR, since the SUR method assumes there are no endogenous regressors. For SUR to be effective, the models must use different regressors. SUR produces the same results as OLS unless the model contains at least one regressor not used in the other equations.

### **Three-Stage Least Squares**

The three-stage least-squares method generalizes the two-stage least-squares method to take account of the correlations between equations in the same way that SUR generalizes OLS. Three-stage least squares requires three steps: first-stage regressions to get predicted values for the endogenous regressors; a two-stage least-squares step to get residuals to estimate the cross-equation correlation matrix; and the final 3SLS estimation step.

The 3SLS option on the PROC SYSLIN statement specifies the three-stage least-squares method, as shown in the following statements.

```
proc syslin data=in 3sls;  
  endogenous p;  
  instruments y u s;  
  demand: model q = p y s;  
  supply: model q = p u;  
run;
```

The 3SLS output begins with a two-stage least-squares regression to estimate the cross-model correlation matrix. This output is the same as the 2SLS results shown in Figure 19.3 and Figure 19.4, and is not repeated here. The next part of the 3SLS output prints the cross-model correlation matrix computed from the 2SLS residuals. This output is shown in Figure 19.5 and includes the cross-model covariances, correlations, the inverse of the correlation matrix, and the inverse covariance matrix.

The SYSLIN Procedure		
Three-Stage Least Squares Estimation		
Cross Model Covariance		
	DEMAND	SUPPLY
DEMAND	0.027892	-.011283
SUPPLY	-.011283	0.018991
Cross Model Correlation		
	DEMAND	SUPPLY
DEMAND	1.00000	-0.49022
SUPPLY	-0.49022	1.00000
Cross Model Inverse Correlation		
	DEMAND	SUPPLY
DEMAND	1.31634	0.64530
SUPPLY	0.64530	1.31634
Cross Model Inverse Covariance		
	DEMAND	SUPPLY
DEMAND	47.1945	28.0380
SUPPLY	28.0380	69.3130

**Figure 19.5.** Estimated Cross-Model Covariances used for 3SLS Estimates

The final 3SLS estimates are shown in Figure 19.6.

```

The SYSLIN Procedure
Three-Stage Least Squares Estimation

System Weighted MSE           0.5711
Degrees of freedom            113
System Weighted R-Square      0.9627

Model                          DEMAND
Dependent Variable             q
Label                          Quantity

Parameter Estimates

Variable      Parameter Standard      Variable
              DF Estimate   Error t Value Pr > |t| Label
Intercept    1  1.980261 1.169169    1.69  0.0959 Intercept
p            1 -1.17654 0.605012   -1.94  0.0568 Price
y            1  0.404115 0.117179    3.45  0.0011 Income
s            1  0.359204 0.085077    4.22 <.0001 Price of Substitutes

Model                          SUPPLY
Dependent Variable             q
Label                          Quantity

Parameter Estimates

Variable      Parameter Standard      Variable
              DF Estimate   Error t Value Pr > |t| Label
Intercept    1  -0.51878 0.490999   -1.06  0.2952 Intercept
p            1  1.333080 0.059271   22.49 <.0001 Price
u            1  -1.14623 0.243491   -4.71 <.0001 Unit Cost
    
```

**Figure 19.6.** Three-Stage Least Squares Results

This output first prints the system weighted mean square error and system weighted  $R^2$  statistics. The system weighted MSE and system weighted  $R^2$  measure the fit of the joint model obtained by stacking all the models together and performing a single regression with the stacked observations weighted by the inverse of the model error variances. See the section "The  $R^2$  Statistics" for details.

Next, the table of 3SLS parameter estimates for each model is printed. This output has the same form as for the other estimation methods.

Note that the 3SLS and 2SLS results may be the same in some cases. This results from the same principle that causes OLS and SUR results to be identical unless an equation includes a regressor not used in the other equations of the system. However, the application of this principle is more complex when instrumental variables are used. When all the exogenous variables are used as instruments, linear combinations of all the exogenous variables appear in the third-stage regressions through substitution of first-stage predicted values.

In this example, 3SLS produces different (and, it is hoped, more efficient) estimates for the demand equation. However, the 3SLS and 2SLS results for the supply equation are the same. This is because the supply equation has one endogenous regressor

and one exogenous regressor not used in other equations. In contrast, the demand equation has fewer endogenous regressors than exogenous regressors not used in other equations in the system.

### Full Information Maximum Likelihood

The FIML option on the PROC SYSLIN statement specifies the full information maximum likelihood method, as shown in the following statements.

```
proc syslin data=in fiml;
  endogenous p q;
  instruments y u s;
  demand: model q = p y s;
  supply: model q = p u;
run;
```

The FIML results are shown in Figure 19.7.

The SYSLIN Procedure									
Full-Information Maximum Likelihood Estimation									
NOTE: Convergence criterion met at iteration 3.									
Model					DEMAND				
Dependent Variable					q				
Label					Quantity				
Parameter Estimates									
Variable	DF	Parameter Estimate	Standard Error	t Value	Pr >  t	Variable	Label		
Intercept	1	1.988529	1.233625	1.61	0.1126	Intercept			
p	1	-1.18147	0.652274	-1.81	0.0755	Price			
y	1	0.402310	0.107269	3.75	0.0004	Income			
s	1	0.361345	0.103816	3.48	0.0010	Price of Substitutes			
Model					SUPPLY				
Dependent Variable					q				
Label					Quantity				
Parameter Estimates									
Variable	DF	Parameter Estimate	Standard Error	t Value	Pr >  t	Variable	Label		
Intercept	1	-0.52443	0.479522	-1.09	0.2787	Intercept			
p	1	1.336083	0.057939	23.06	<.0001	Price			
u	1	-1.14804	0.237793	-4.83	<.0001	Unit Cost			

Figure 19.7. FIML Results

## Computing Reduced Form Estimates

A system of structural equations with endogenous regressors can be represented as functions only of the predetermined variables. For this to be possible, there must be

Part 2. General Information

as many equations as endogenous variables. If there are more endogenous variables than regression models, you can use IDENTITY statements to complete the system. See "Reduced Form Estimates" in the "Computational Details" section later in this chapter for details.

The REDUCED option on the PROC SYSLIN statement prints reduced form estimates. The following statements show this using the 3SLS estimates of the structural parameters.

```
proc syslin data=in 3sls reduced;
  endogenous p;
  instruments y u s;
  demand: model q = p y s;
  supply: model q = p u;
run;
```

The first four pages of this output were as shown previously and are not repeated here. (See Figure 19.3, Figure 19.4, Figure 19.5, and Figure 19.6.) The final page of the output from this example contains the reduced form coefficients from the 3SLS structural estimates, as shown in Figure 19.8.

The SYSLIN Procedure				
Three-Stage Least Squares Estimation				
Endogenous Variables				
		p	q	
DEMAND		1.176539	1	
SUPPLY		-1.33308	1	
Exogenous Variables				
	Intercept	y	s	u
DEMAND	1.980261	0.404115	0.359204	0
SUPPLY	-0.51878	0	0	-1.14623
Inverse Endogenous Variables				
		DEMAND	SUPPLY	
p		0.398467	-0.39847	
q		0.531188	0.468812	
Reduced Form				
	Intercept	y	s	u
p	0.995786	0.161027	0.143131	0.456736
q	0.80868	0.214661	0.190805	-0.53737

Figure 19.8. Reduced Form 3SLS Results

---

## Restricting Parameter Estimates

You can impose restrictions on the parameter estimates with `RESTRICT` and `SRESTRICT` statements. The `RESTRICT` statement imposes linear restrictions on parameters in the equation specified by the preceding `MODEL` statement. The `SRESTRICT` statement imposes linear restrictions that relate parameters in different models.

To impose restrictions involving parameters in different equations, use the `SRESTRICT` statement. Specify the parameters in the linear hypothesis as *model-label.regressor-name*. (If the `MODEL` statement does not have a label, you can use the dependent variable name as the label for the model, provided the dependent variable uniquely labels the model.)

Tests for the significance of the restrictions are printed when `RESTRICT` or `SRESTRICT` statements are used. You can label `RESTRICT` and `SRESTRICT` statements to identify the restrictions in the output.

The `RESTRICT` statement in the following example restricts the price coefficient in the demand equation to equal .015. The `SRESTRICT` statement restricts the estimate of the income coefficient in the demand equation to be .01 times the estimate of the unit cost coefficient in the supply equation.

```
proc syslin data=in 3sls;
  endogenous p;
  instruments y u s;
  demand: model q = p y s;
  peq015: restrict p = .015;
  supply: model q = p u;
  yeq01u: srestring demand.y = .01 * supply.u;
run;
```

The restricted estimation results are shown in Figure 19.9.

```

The SYSLIN Procedure
Three-Stage Least Squares Estimation

Model DEMAND
Dependent Variable q
Label Quantity

Parameter Estimates
Variable DF Parameter Estimate Standard Error t Value Pr > |t| Variable Label
Intercept 1 -0.46584 0.053307 -8.74 <.0001 Intercept
p 1 0.015000 0 . . Price
y 1 -0.00679 0.002357 -2.88 0.0056 Income
s 1 0.325589 0.009872 32.98 <.0001 Price of Substitutes
RESTRICT -1 50.59341 7.464990 6.78 <.0001 PEQ015

Model SUPPLY
Dependent Variable q
Label Quantity

Parameter Estimates
Variable DF Parameter Estimate Standard Error t Value Pr > |t| Variable Label
Intercept 1 -1.31894 0.477633 -2.76 0.0077 Intercept
p 1 1.291718 0.059101 21.86 <.0001 Price
u 1 -0.67887 0.235679 -2.88 0.0056 Unit Cost

Parameter Estimates
Variable DF Parameter Estimate Standard Error t Value Pr > |t| Variable Label
RESTRICT -1 342.3611 38.12103 8.98 <.0001 YEQ01U

```

**Figure 19.9.** Restricted Estimates

The standard error for P in the demand equation is 0, since the value of the P coefficient was specified by the RESTRICT statement and not estimated from the data. The Parameter Estimates table for the demand equation contains an additional row for the restriction specified by the RESTRICT statement. The "parameter estimate" for the restriction is the value of the Lagrange multiplier used to impose the restriction. The restriction is highly "significant" ( $t = 6.777$ ), which means that the data are not consistent with the restriction, and the model does not fit as well with the restriction imposed. See the section "RESTRICT Statement" for more information.

After the Parameter Estimates table for the supply equation, the results for the cross model restrictions are printed. This shows that the restriction specified by the SRESTRICT statement is not consistent with the data ( $t = 8.98$ ). See the section "SRESTRICT Statement" for more information.



## Testing Parameters

You can test linear hypotheses about the model parameters with TEST and STEST statements. The TEST statement tests hypotheses about parameters in the equation specified by the preceding MODEL statement. The STEST statement tests hypotheses that relate parameters in different models.

For example, the following statements test the hypothesis that the price coefficient in the demand equation is equal to .015.

```
proc syslin data=in 3sls;
  endogenous p;
  instruments y u s;
  demand: model q = p y s;
  test_1: test p = .015;
  supply: model q = p u;
run;
```

The TEST statement results are shown in Figure 19.10. This reports an  $F$ -test for the hypothesis specified by the TEST statement. In this case the  $F$  statistic is 6.79 (3.879/.571) with 1 and 113 degrees of freedom. The  $p$ -value for this  $F$  statistic is .0104, which indicates that the hypothesis tested is almost but not quite rejected at the .01 level. See the section "TEST Statement" for more information.

The SYSLIN Procedure						
Three-Stage Least Squares Estimation						
System Weighted MSE		0.5711				
Degrees of freedom		113				
System Weighted R-Square		0.9627				
Model		DEMAND				
Dependent Variable		q				
Label		Quantity				
Parameter Estimates						
Variable	DF	Parameter Estimate	Standard Error	t Value	Pr >  t	Variable Label
Intercept	1	1.980261	1.169169	1.69	0.0959	Intercept
p	1	-1.17654	0.605012	-1.94	0.0568	Price
y	1	0.404115	0.117179	3.45	0.0011	Income
s	1	0.359204	0.085077	4.22	<.0001	Price of Substitutes
Test Results for Variable TEST_1						
Num DF	Den DF	F Value	Pr > F			
1	113	6.79	0.0104			

Figure 19.10. TEST Statement Results

## Part 2. General Information

To test hypotheses involving parameters in different equations, use the STEST statement. Specify the parameters in the linear hypothesis as *model-label.regressor-name*. (If the MODEL statement does not have a label, you can use the dependent variable name as the label for the model, provided the dependent variable uniquely labels the model.)

For example, the following statements test the hypothesis that the income coefficient in the demand equation is .01 times the unit cost coefficient in the supply equation:

```
proc syslin data=in 3sls;
  endogenous p;
  instruments y u s;
  demand: model q = p y s;
  supply: model q = p u;
  stest1: stest demand.y = .01 * supply.u;
run;
```

The STEST statement results are shown in Figure 19.11. The form and interpretation of the STEST statement results is like the TEST statement results. In this case, the *F*-test produces a *p*-value less than .0001, and strongly rejects the hypothesis tested. See the section "STEST Statement" for more information.

```

The SYSLIN Procedure
Three-Stage Least Squares Estimation

System Weighted MSE           0.5711
Degrees of freedom            113
System Weighted R-Square      0.9627

Model                          DEMAND
Dependent Variable             q
Label                          Quantity

Parameter Estimates

Variable      Parameter Standard      Variable
              DF Estimate      Error t Value Pr > |t| Label
Intercept    1  1.980261 1.169169    1.69  0.0959 Intercept
p            1 -1.17654 0.605012   -1.94  0.0568 Price
y            1  0.404115 0.117179    3.45  0.0011 Income
s            1  0.359204 0.085077    4.22  <.0001 Price of Substitutes

Model                          SUPPLY
Dependent Variable             q
Label                          Quantity

Parameter Estimates

Variable      Parameter Standard      Variable
              DF Estimate      Error t Value Pr > |t| Label
Intercept    1 -0.51878 0.490999   -1.06  0.2952 Intercept
p            1  1.333080 0.059271   22.49  <.0001 Price
u            1 -1.14623 0.243491   -4.71  <.0001 Unit Cost

Test Results for Variable STEST1

              Num DF      Den DF      F Value      Pr > F
              1          113          22.46      0.0001

```

**Figure 19.11.** STEST Statement Results

You can combine TEST and STEST statements with RESTRICT and SRESTRICT statements to perform hypothesis tests for restricted models. Of course, the validity of the TEST and STEST statement results will depend on the correctness of any restrictions you impose on the estimates.

---

## Saving Residuals and Predicted Values

You can store predicted values and residuals from the estimated models in a SAS data set. Specify the `OUT=` option on the `PROC SYSLIN` statement and use the `OUTPUT` statement to specify names for new variables to contain the predicted and residual values.

For example, the following statements store the predicted quantity from the supply and demand equations in a data set `PRED`:

```
proc syslin data=in out=pred 3sls;
  endogenous p;
  instruments y u s;
  demand: model q = p y s;
  output predicted=q_demand;
  supply: model q = p u;
  output predicted=q_supply;
run;
```

---

## Plotting Residuals

You can plot the residuals against the regressors by specifying the `PLOT` option on the `MODEL` statement. For example, the following statements plot the 2SLS residuals for the demand model against price, income, price of substitutes, and the intercept.

```
proc syslin data=in 2sls;
  endogenous p;
  instruments y u s;
  demand: model q = p y s / plot;
run;
```

The plot for price is shown in Figure 19.12. The other plots are not shown.

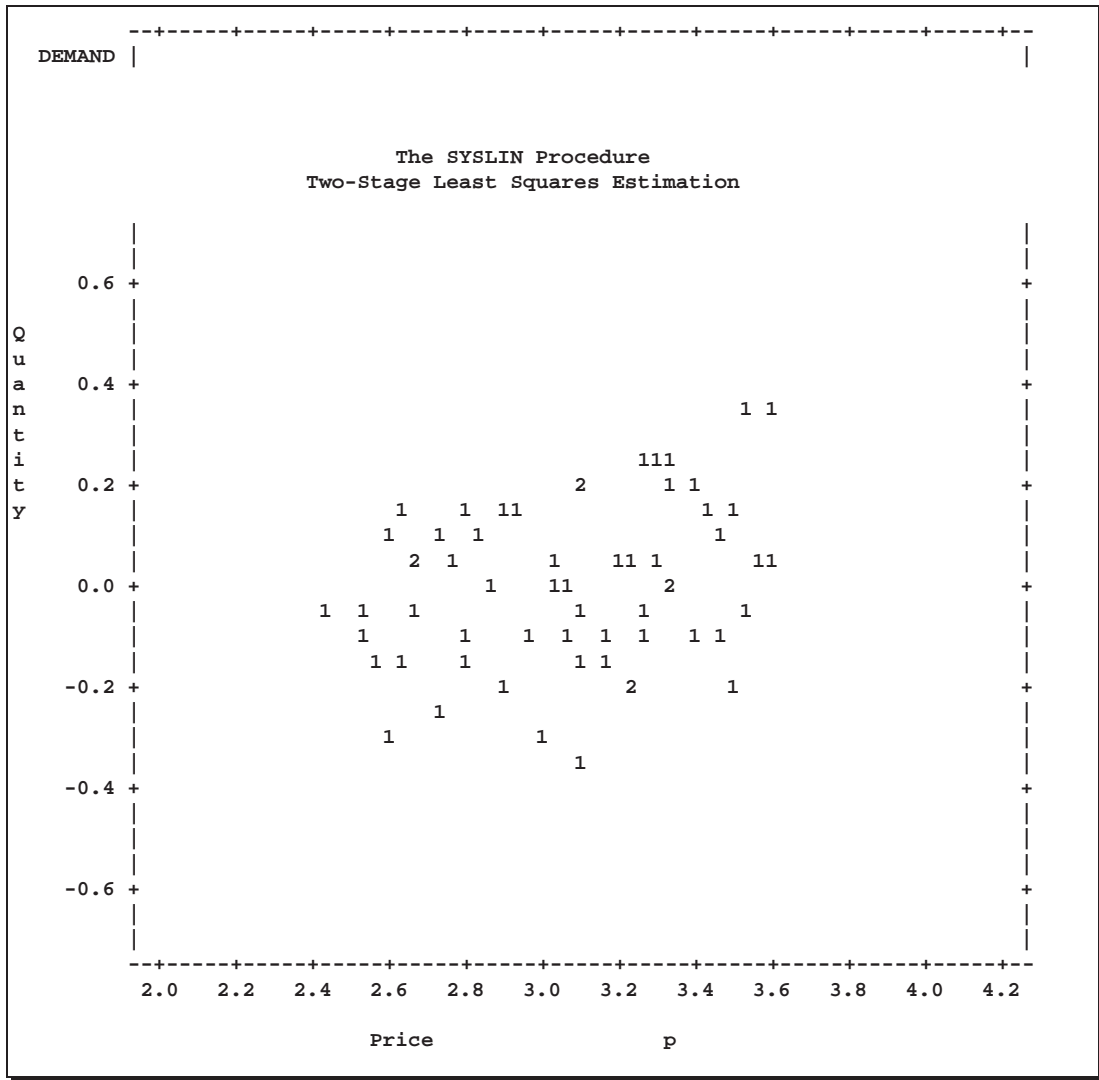


Figure 19.12. PLOT Option Output for P

---

## Syntax

The SYSLIN procedure uses the following statements:

```

PROC SYSLIN options ;
  BY variables ;
  ENDOGENOUS variables ;
  IDENTITY identities ;
  INSTRUMENTS variables ;
  MODEL response = regressors / options ;
  OUTPUT PREDICTED= variable RESIDUAL= variable ;
  RESTRICT restrictions ;
  SRESTRICT restrictions ;
  STEST equations ;
  TEST equations ;
  VAR variables ;
  WEIGHT variable ;

```

---

## Functional Summary

The SYSLIN procedure statements and options are summarized in the following table.

Description	Statement	Option
<b>Data Set Options</b>		
specify the input data set	PROC SYSLIN	DATA=
specify the output data set	PROC SYSLIN	OUT=
write parameter estimates to an output data set	PROC SYSLIN	OUTEST=
write covariances to the OUTEST= data set	PROC SYSLIN	OUTCOV OUTCOV3
write the SSCP matrix to an output data set	PROC SYSLIN	OUTSSCP=
<b>Estimation Method Options</b>		
specify full information maximum likelihood estimation	PROC SYSLIN	FIML
specify iterative SUR estimation	PROC SYSLIN	ITSUR
specify iterative 3SLS estimation	PROC SYSLIN	IT3SLS
specify K-class estimation	PROC SYSLIN	K=
specify limited information maximum likelihood estimation	PROC SYSLIN	LIML
specify minimum expected loss estimation	PROC SYSLIN	MELO
specify ordinary least squares estimation	PROC SYSLIN	OLS
specify seemingly unrelated estimation	PROC SYSLIN	SUR
specify two-stage least-squares estimation	PROC SYSLIN	2SLS

Description	Statement	Option
specify three-stage least-squares estimation	PROC SYSLIN	3SLS
specify Fuller's modification to LIML	PROC SYSLIN	ALPHA=
specify convergence criterion	PROC SYSLIN	CONVERGE=
specify maximum number of iterations	PROC SYSLIN	MAXIT=
use diagonal of <b>S</b> instead of <b>S</b>	PROC SYSLIN	SDIAG
exclude RESTRICT statements in final stage	PROC SYSLIN	NOINCLUDE
specify criterion for testing for singularity	PROC SYSLIN	SINGULAR=
specify denominator for variance estimates	PROC SYSLIN	VARDEF=
<b>Printing Control Options</b>		
print first-stage regression statistics	PROC SYSLIN	FIRST
print estimates and SSE at each iteration	PROC SYSLIN	ITPRINT
print the restricted reduced form estimates	PROC SYSLIN	REDUCED
print descriptive statistics	PROC SYSLIN	SIMPLE
print uncorrected SSCP matrix	PROC SYSLIN	USSCP
print correlations of the parameter estimates	MODEL	CORRB
print covariances of the parameter estimates	MODEL	COVB
print Durbin-Watson statistics	MODEL	DW
print Basman's test	MODEL	OVERID
plot residual values against regressors	MODEL	PLOT
print standardized parameter estimates	MODEL	STB
print unrestricted parameter estimates	MODEL	UNREST
print the model crossproducts matrix	MODEL	XPX
print the inverse of the crossproducts matrix	MODEL	I
suppress printed output	MODEL	NOPRINT
suppress all printed output	PROC SYSLIN	NOPRINT
<b>Model Specification</b>		
specify structural equations	MODEL	
suppress the intercept parameter	MODEL	NOINT
specify linear relationship among variables	IDENTITY	
perform weighted regression	WEIGHT	
<b>Tests and Restrictions on Parameters</b>		
place restrictions on parameter estimates	RESTRICT	
place restrictions on parameter estimates	SRESTRICT	
test linear hypothesis	STEST	
test linear hypothesis	TEST	
<b>Other Statements</b>		
specify BY-group processing	BY	

Description	Statement	Option
specify the endogenous variables	ENDOGENOUS	
specify instrumental variables	INSTRUMENTS	
write predicted and residual values to a data set	OUTPUT	
name variable for predicted values	OUTPUT	PREDICTED=
name variable for residual values	OUTPUT	RESIDUAL=
include additional variables in $X'X$ matrix	VAR	

## PROC SYSLIN Statement

### PROC SYSLIN *options*;

The following options can be used with the PROC SYSLIN statement.

#### Data Set Options

##### DATA= *SAS-data-set*

specifies the input data set. If the DATA= option is omitted, the most recently created SAS data set is used. In addition to ordinary SAS data sets, PROC SYSLIN can analyze data sets of TYPE=CORR, TYPE=COV, TYPE=UCORR, TYPE=UCOV, and TYPE=SSCP. See "Special TYPE= Input Data Set" in the "Input Data Set" section later in this chapter for more information.

##### OUT= *SAS-data-set*

specifies an output SAS data set for residuals and predicted values. The OUT= option is used in conjunction with the OUTPUT statement. See the section "OUT= Data Set" later in this chapter for more details.

##### OUTEST= *SAS-data-set*

writes the parameter estimates to an output data set. See the section "OUTEST= Data Set" later in this chapter for details.

##### OUTCOV

##### COVOUT

writes the covariance matrix of the parameter estimates to the OUTEST= data set in addition to the parameter estimates.

##### OUTCOV3

##### COV3OUT

writes covariance matrices for each model in a system to the OUTEST= data set when the 3SLS, SUR, or FIML option is used.

##### OUTSSCP= *SAS-data-set*

writes the sum-of-squares-and-crossproducts matrix to an output data set. See the section "OUTSSCP= Data Set" later in this chapter for details.



**Estimation Method Options****2SLS**

specifies the two-stage least-squares estimation method.

**3SLS**

specifies the three-stage least-squares estimation method.

**FIML**

specifies the full information maximum likelihood estimation method.

**ITSUR**

specifies the iterative seemingly unrelated estimation method.

**IT3SLS**

specifies the iterative three-stage least-squares estimation method.

**K= *value***

specifies the K-class estimation method.

**LIML**

specifies the limited information maximum likelihood estimation method.

**MELO**

specifies the minimum expected loss estimation method.

**OLS**

specifies the ordinary least squares estimation method. This is the default.

**SUR**

specifies the seemingly unrelated estimation method.

**Printing and Control Options****ALL**

specifies the CORRB, COVB, DW, I, OVERID, PLOT, STB, and XPX options for every MODEL statement.

**ALPHA= *value***

specifies Fuller's modification to the LIML estimation method. See "Fuller's Modification to LIML K Value" later in this chapter for details.

**CONVERGE= *value***

specifies the convergence criterion for the iterative estimation methods IT3SLS, ITSUR, and FIML. The default is CONVERGE=.0001.

**FIRST**

prints first-stage regression statistics for the endogenous variables regressed on the instruments. This output includes sums of squares, estimates, variances, and standard deviations.

**ITPRINT**

prints parameter estimates, system-weighted residual sum of squares, and  $R^2$  at each iteration for the IT3SLS and ITSUR estimation methods. For the FIML method, the ITPRINT option prints parameter estimates, negative of log likelihood function, and norm of gradient vector at each iteration.

**MAXITER= *n***

specifies the maximum number of iterations allowed for the IT3SLS, ITSUR, and FIML estimation methods. The MAXITER= option can be abbreviated as MAXIT=. The default is MAXITER=30.

**NOINCLUDE**

excludes the RESTRICT statements from the final stage for the 3SLS, IT3SLS, SUR, ITSUR estimation methods.

**NOPRINT**

suppresses all printed output. Specifying NOPRINT in the PROC SYSLIN statement is equivalent to specifying NOPRINT in every MODEL statement.

**REDUCED**

prints the reduced form estimates. If the REDUCED option is specified, you should specify any IDENTITY statements needed to make the system square. See "Reduced Form Estimates" in the section "Computational Details" later in this chapter for more information.

**SDIAG**

uses the diagonal of **S** instead of **S** to do the estimation, where **S** is the covariance matrix of equation errors. See "Uncorrelated Errors Across Equations" in the section "Computational Details" later in this chapter for more information.

**SIMPLE**

prints descriptive statistics for the dependent variables. The statistics printed include the sum, mean, uncorrected sum of squares, variance, and standard deviation.

**SINGULAR= *value***

specifies a criterion for testing singularity of the crossproducts matrix. This is a tuning parameter used to make PROC SYSLIN more or less sensitive to singularities. The value must be between 0 and 1. The default is SINGULAR=1E-8.

**USSCP**

prints the uncorrected sum-of-squares-and-crossproducts matrix.

**USSCP2**

prints the uncorrected sum-of-squares-and-crossproducts matrix for all variables used in the analysis, including predicted values of variables generated by the procedure.

**VARDEF= DF | N | WEIGHT | WGT**

specifies the denominator to use in calculating cross-equation error covariances and parameter standard errors and covariances. The default is VARDEF=DF, which corrects for model degrees of freedom. VARDEF=N specifies no degrees-of-freedom correction. VARDEF=WEIGHT specifies the sum of the observation weights. VARDEF=WGT specifies the sum of the observation weights minus the model degrees of freedom. See "Computation of Standard Errors" in the section "Computational Details" later in this chapter for more information.

---

## BY Statement

**BY** *variables* ;

A BY statement can be used with PROC SYSLIN to obtain separate analyses on observations in groups defined by the BY variables.

---

## ENDOGENOUS Statement

**ENDOGENOUS** *variables* ;

The ENDOGENOUS statement declares the jointly dependent variables that are projected in the first-stage regression through the instrument variables. The ENDOGENOUS statement is not needed for the SUR, ITSUR, or OLS estimation methods. The default ENDOGENOUS list consists of all the dependent variables in the MODEL and IDENTITY statements that do not appear in the INSTRUMENTS statement.

---

## IDENTITY Statement

**IDENTITY** *equation* ;

The IDENTITY statement specifies linear relationships among variables to write to the OUTEST= data set. It provides extra information in the OUTEST= data set but does not create or compute variables. The OUTEST= data set can be processed by the SIMLIN procedure in a later step.

The IDENTITY statement is also used to compute reducedform coefficients when the REDUCED option in the PROC SYSLIN statement is specified. See "Reduced Form Estimates" in the section "Computational Details" later in this chapter for more information.

The *equation* given by the IDENTITY statement has the same form as equations in the MODEL statement. A label can be specified for an IDENTITY statement as follows:

*label*: **IDENTITY** ... ;

---

## INSTRUMENTS Statement

**INSTRUMENTS** *variables* ;

The INSTRUMENTS statement declares the variables used in obtaining first-stage predicted values. All the instruments specified are used in each first-stage regression. The INSTRUMENTS statement is required for the 2SLS, 3SLS, IT3SLS,

LIML, MELO, and K-class estimation methods. The INSTRUMENTS statement is not needed for the SUR, ITSUR, OLS, or FIML estimation methods.

---

## MODEL Statement

**MODEL** *response = regressors / options ;*

The MODEL statement regresses the response variable on the left side of the equal sign against the regressors listed on the right side.

Models can be given labels. Model labels are used in the printed output to identify the results for different models. Model labels are also used in SRESTRICT and STEST statements to refer to parameters in different models. If no label is specified, the response variable name is used as the label for the model. The model label is specified as follows:

*label:* **MODEL** . . . ;

The following options can be used in the MODEL statement after a slash (/).

### **ALL**

specifies the CORRB, COVB, DW, I, OVERID, PLOT, STB, and XPX options.

### **ALPHA=** *value*

specifies the  $\alpha$  parameter for Fuller's modification to the LIML estimation method. See "Fuller's Modification to LIML" in the section "Computational Details" later in this chapter for more information.

### **CORRB**

prints the matrix of estimated correlations between the parameter estimates.

### **COVB**

prints the matrix of estimated covariances between the parameter estimates.

### **DW**

prints Durbin-Watson statistics and autocorrelation coefficients for the residuals. If there are missing values,  $d'$  is calculated according to Savin and White (1978). Use the DW option only if the data set to be analyzed is an ordinary SAS data set with time series observations sorted in time order. The Durbin-Watson test is not valid for models with lagged dependent regressors.

### **I**

prints the inverse of the crossproducts matrix for the model,  $(\mathbf{X}'\mathbf{X})^{-1}$ . If restrictions are specified, the crossproducts matrix printed is adjusted for the restrictions. See the section "Computational Details" for more information.

### **K=** *value*

specifies K-class estimation.

### **NOINT**

suppresses the intercept parameter from the model.

**NOPRINT**

suppresses the normal printed output.

**OVERID**

prints Basman's (1960) test for over identifying restrictions. See "Over Identification Restrictions" in the section "Computational Details" later in this chapter for more information.

**PLOT**

plots residual values against regressors. A plot of the residuals for each regressor is printed.

**STB**

prints standardized parameter estimates. Sometimes known as a standard partial regression coefficient, a standardized parameter estimate is a parameter estimate multiplied by the standard deviation of the associated regressor and divided by the standard deviation of the response variable.

**UNREST**

prints parameter estimates computed before restrictions are applied. The UNREST option is valid only if a RESTRICT statement is specified.

**XPX**

prints the model crossproducts matrix,  $X'X$ . See the section "Computational Details" for more information.

---

## OUTPUT Statement

**OUTPUT PREDICTED=***variable* **RESIDUAL=***variable* ;

The OUTPUT statement writes predicted values and residuals from the preceding model to the data set specified by the OUT= option on the PROC SYSLIN statement. An OUTPUT statement must come after the MODEL statement to which it applies. The OUT= option must be specified in the PROC SYSLIN statement.

The following options can be specified in the OUTPUT statement:

**PREDICTED=** *variable*

names a new variable to contain the predicted values for the response variable. The PREDICTED= option can be abbreviated as PREDICT=, PRED=, or P=.

**RESIDUAL=** *variable*

names a new variable to contain the residual values for the response variable. The RESIDUAL= option can be abbreviated as RESID= or R=.

For example, the following statements create an output data set named B. In addition to the variables in the input data set, the data set B contains the variable YHAT, with values that are predicted values of the response variable Y, and YRESID, with values that are the residual values of Y.

```
proc syslin data=a out=b;  
  model y = x1 x2;  
  output p=yhat r=yresid;  
run;
```

For example, the following statements create an output data set named PRED. In addition to the variables in the input data set, the data set PRED contains the variables Q\_DEMAND and Q\_SUPPLY, with values that are predicted values of the response variable Q for the demand and supply equations respectively, and R\_DEMAND and R\_SUPPLY, with values that are the residual values of the demand and supply equations.

```
proc syslin data=in out=pred;  
  demand: model q = p y s;  
  output p=q_demand r=r_demand;  
  supply: model q = p u;  
  output p=q_supply r=r_supply;  
run;
```

See the section "OUT= Data Set" later in this chapter for more details.

---

## RESTRICT Statement

**RESTRICT** *equation* , ... , *equation* ;

The RESTRICT statement places restrictions on the parameter estimates for the preceding MODEL statement. Any number of restrict statements can follow a MODEL statement. Each restriction is written as a linear equation. If more than one restriction is specified in a single RESTRICT statement, the restrictions are separated by commas.

Parameters are referred to by the name of the corresponding regressor variable. Each name used in the equation must be a regressor in the preceding MODEL statement. The keyword INTERCEPT is used to refer to the intercept parameter in the model.

RESTRICT statements can be given labels. The labels are used in the printed output to distinguish results for different restrictions. Labels are specified as follows:

*label* : **RESTRICT** ... ;

The following is an example of the use of the RESTRICT statement, in which the coefficients of the regressors X1 and X2 are required to sum to 1.

```
proc syslin data=a;  
  model y = x1 x2;  
  restrict x1 + x2 = 1;  
run;
```

Variable names can be multiplied by constants. When no equal sign appears, the linear combination is set equal to 0. Note that the parameters associated with the variables are restricted, not the variables themselves. Here are some examples of valid RESTRICT statements:

```
restrict x1 + x2 = 1;
restrict x1 + x2 - 1;
restrict 2 * x1 = x2 + x3 , intercept + x4 = 0;
restrict x1 = x2 = x3 = 1;
restrict 2 * x1 - x2;
```

Restricted parameter estimates are computed by introducing a Lagrangian parameter  $\lambda$  for each restriction (Pringle and Raynor 1971). The estimates of these Lagrangian parameters are printed in the parameter estimates table. If a restriction cannot be applied, its parameter value and degrees of freedom are listed as 0.

The Lagrangian parameter,  $\lambda$ , measures the sensitivity of the SSE to the restriction. If the restriction is changed by a small amount  $\epsilon$ , the SSE is changed by  $2\lambda\epsilon$ .

The  $t$ -ratio tests the significance of the restrictions. If  $\lambda$  is zero, the restricted estimates are the same as the unrestricted.

Any number of restrictions can be specified on a RESTRICT statement, and any number of RESTRICT statements can be used. The estimates are computed subject to all restrictions specified. However, restrictions should be consistent and not redundant.

Note: The RESTRICT statement is not supported for the FIML estimation method.

---

## SRESTRICT Statement

**SRESTRICT** *equation* , ... , *equation* ;

The SRESTRICT statement imposes linear restrictions involving parameters in two or more MODEL statements. The SRESTRICT statement is like the RESTRICT statement but is used to impose restrictions across equations, whereas the RESTRICT statement only applies to parameters in the immediately preceding MODEL statement.

Each restriction is written as a linear equation. Parameters are referred to as *label.variable*, where *label* is the model label and *variable* is the name of the regressor to which the parameter is attached. (If the MODEL statement does not have a label, you can use the dependent variable name as the label for the model, provided the dependent variable uniquely labels the model.) Each variable name used must be a regressor in the indicated MODEL statement. The keyword INTERCEPT is used to refer to intercept parameters.

SRESTRICT statements can be given labels. The labels are used in the printed output to distinguish results for different restrictions. Labels are specified as follows:

*label* : **SRESTRICT** ... ;

## Part 2. General Information

The following is an example of the use of the SRESTRICT statement, in which the coefficient for the regressor X2 is constrained to be the same in both models.

```
proc syslin data=a 3sls;
  endogenous y1 y2;
  instruments x1 x2;
  model y1 = y2 x1 x2;
  model y2 = y1 x2;
  srestrict y1.x2 = y2.x2;
run;
```

When no equal sign is used, the linear combination is set equal to 0. Thus the restriction in the preceding example can also be specified as

```
srestrict y1.x2 - y2.x2;
```

Any number of restrictions can be specified on an SRESTRICT statement, and any number of SRESTRICT statements can be used. The estimates are computed subject to all restrictions specified. However, restrictions should be consistent and not redundant.

The results of the SRESTRICT statements are printed after the parameter estimates for all the models in the system. The format of the SRESTRICT statement output is the same as the parameter estimates table. In this output the "Parameter Estimate" is the Lagrangian parameter,  $\lambda$ , used to impose the restriction.

The Lagrangian parameter,  $\lambda$ , measures the sensitivity of the system sum of square errors to the restriction. The system SSE is the system MSE shown in the printed output multiplied by the degrees of freedom. If the restriction is changed by a small amount  $\epsilon$ , the system SSE is changed by  $2\lambda\epsilon$ .

The *t*-ratio tests the significance of the restriction. If  $\lambda$  is zero, the restricted estimates are the same as the unrestricted estimates.

The model degrees of freedom are not adjusted for the cross-model restrictions imposed by SRESTRICT statements.

Note: The SRESTRICT statement is not supported for the FIML estimation method.

---

## STEST Statement

**STEST** *equation , ... , equation / options ;*

The STEST statement performs an *F*-test for the joint hypotheses specified in the statement.

The hypothesis is represented in matrix notation as

$$\mathbf{L}\beta = \mathbf{c}$$



and the  $F$ -test is computed as

$$\frac{(\mathbf{L}b - \mathbf{c})'(\mathbf{L}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{L}')^{-1}(\mathbf{L}b - \mathbf{c})}{m\hat{\sigma}^2}$$

where  $b$  is the estimate of  $\beta$ ,  $m$  is the number of restrictions, and  $\hat{\sigma}^2$  is the system weighted mean square error. See the section "Computational Details" for information on the matrix  $\mathbf{X}'\mathbf{X}$ .

Each hypothesis to be tested is written as a linear equation. Parameters are referred to as *label.variable*, where *label* is the model label and *variable* is the name of the regressor to which the parameter is attached. (If the MODEL statement does not have a label, you can use the dependent variable name as the label for the model, provided the dependent variable uniquely labels the model.) Each variable name used must be a regressor in the indicated MODEL statement. The keyword INTERCEPT is used to refer to intercept parameters.

STEST statements can be given labels. The label is used in the printed output to distinguish different tests. Any number of STEST statements can be specified. Labels are specified as follows:

*label*: **STEST** ... ;

The following is an example of the STEST statement:

```
proc syslin data=a 3sls;
  endogenous y1 y2;
  instruments x1 x2;
  model y1 = y2 x1 x2;
  model y2 = y1 x2;
  stest y1.x2 = y2.x2;
run;
```

The test performed is exact only for ordinary least squares, given the OLS assumptions of the linear model. For other estimation methods, the  $F$ -test is based on large sample theory and is only approximate in finite samples.

If RESTRICT or SRESTRICT statements are used, the tests computed by the STEST statement are conditional on the restrictions specified. The validity of the tests may be compromised if incorrect restrictions are imposed on the estimates.

The following are examples of STEST statements:

```
stest a.x1 + b.x2 = 1;
stest 2 * b.x2 = c.x3 + c.x4 ,
      a.intercept + b.x2 = 0;
stest a.x1 = c.x2 = b.x3 = 1;
stest 2 * a.x1 - b.x2 = 0;
```

The PRINT option can be specified in the STEST statement after a slash (/):

**PRINT**

prints intermediate calculations for the hypothesis tests.

Note: The STEST statement is not supported for the FIML estimation method.

---

## TEST Statement

**TEST** *equation* , ... , *equation* / *options* ;

The TEST statement performs *F*-tests of linear hypotheses about the parameters in the preceding MODEL statement. Each equation specifies a linear hypothesis to be tested. If more than one equation is specified, the equations are separated by commas.

Variable names must correspond to regressors in the preceding MODEL statement, and each name represents the coefficient of the corresponding regressor. The keyword INTERCEPT is used to refer to the model intercept.

TEST statements can be given labels. The label is used in the printed output to distinguish different tests. Any number of TEST statements can be specified. Labels are specified as follows:

*label*: **TEST** ... ;

The following is an example of the use of TEST statement, which tests the hypothesis that the coefficients of X1 and X2 are the same:

```
proc syslin data=a;
  model y = x1 x2;
  test x1 = x2;
run;
```

The following statements perform *F*-tests for the hypothesis that the coefficients of X1 and X2 are equal, and that the sum of the X1 and X2 coefficients is twice the intercept, and for the joint hypothesis.

```
proc syslin data=a;
  model y = x1 x2;
  x1eqx2: test x1 = x2;
  sumeq2i: test x1 + x2 = 2 * intercept;
  joint: test x1 = x2, x1 + x2 = 2 * intercept;
run;
```

The following are additional examples of TEST statements:

```
test x1 + x2 = 1;
test x1 = x2 = x3 = 1;
test 2 * x1 = x2 + x3, intercept + x4 = 0;
test 2 * x1 - x2;
```

The TEST statement performs an  $F$ -test for the joint hypotheses specified. The hypothesis is represented in matrix notation as follows:

$$\mathbf{L}\beta = \mathbf{c}$$

The  $F$  test is computed as

$$\frac{(\mathbf{L}b - \mathbf{c})'(\mathbf{L}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{L}')^{-1}(\mathbf{L}b - \mathbf{c})}{m\hat{\sigma}^2}$$

where  $b$  is the estimate of  $\beta$ ,  $m$  is the number of restrictions, and  $\hat{\sigma}^2$  is the model mean square error. See the section "Computational Details" for information on the matrix  $\mathbf{X}'\mathbf{X}$ .

The test performed is exact only for ordinary least squares, given the OLS assumptions of the linear model. For other estimation methods, the  $F$ -test is based on large sample theory and is only approximate in finite samples.

If RESTRICT or SRESTRICT statements are used, the tests computed by the TEST statement are conditional on the restrictions specified. The validity of the tests may be compromised if incorrect restrictions are imposed on the estimates.

The PRINT option can be specified in the TEST statement after a slash (/):

#### PRINT

prints intermediate calculations for the hypothesis tests.

Note: The TEST statement is not supported for the FIML estimation method.

---

## VAR Statement

**VAR** *variables* ;

The VAR statement is used to include variables in the crossproducts matrix that are not specified in any MODEL statement. This statement is rarely used with PROC SYSLIN and is used only with the OUTSSCP= option in the PROC SYSLIN statement.

---

## WEIGHT Statement

**WEIGHT** *variable* ;

The WEIGHT statement is used to perform weighted regression. The WEIGHT statement names a variable in the input data set whose values are relative weights for a weighted least-squares fit. If the weight value is proportional to the reciprocal of the variance for each observation, the weighted estimates are the best linear unbiased estimates (BLUE).

---

## Details

---

### Input Data Set

PROC SYSLIN does not compute new values for regressors. For example, if you need a lagged variable, you must create it with a DATA step. No values are computed by IDENTITY statements; all values must be in the input data set.

#### **Special TYPE= Input Data Set**

The input data set for most applications of the SYSLIN procedure contains standard rectangular data. However, PROC SYSLIN can also process input data in the form of a crossproducts, covariance, or correlation matrix. Data sets containing such matrices are identified by values of the TYPE= data set option.

These special kinds of input data sets can be used to save computer time. It takes  $nk^2$  operations, where  $n$  is the number of observations and  $k$  is the number of variables, to calculate cross products; the regressions are of the order  $k^3$ . When  $n$  is in the thousands and  $k$  is much smaller, you can save most of the computer time in later runs of PROC SYSLIN by re-using the SSCP matrix rather than recomputing it.

The SYSLIN procedure can process TYPE= CORR, COV, UCORR, UCOV, or SSCP data sets. TYPE=CORR and TYPE=COV data sets, usually created by the CORR procedure, contain means and standard deviations, and correlations or covariances. TYPE=SSCP data sets, usually created in previous runs of PROC SYSLIN, contain sums of squares and cross products. Refer to *SAS/STAT User's Guide* for more information on special SAS data sets.

When special SAS data sets are read, you must specify the TYPE= data set option. PROC CORR and PROC SYSLIN automatically set the type for output data sets; however, if you create the data set by some other means, you must specify its type with the TYPE= data set option.

When the special data sets are used, the DW (Durbin-Watson test) and PLOT options in the MODEL statement cannot be performed, and the OUTPUT statements are not valid.

---

### Estimation Methods

A brief description of the methods used by the SYSLIN procedure follows. For more information on these methods, see the references at the end of this chapter.

There are two fundamental methods of estimation for simultaneous equations: least squares and maximum likelihood. There are two approaches within each of these categories: single equation methods and system estimation. 2SLS, 3SLS, and IT3SLS use the least-squares method; LIML and FIML use the maximum likelihood method. 2SLS and LIML are single equation methods, which means that over identifying restrictions in other equations are not taken into account in estimating parameters in a particular equation. (See "Over Identification Restrictions" in the section "Computational Details" later in this chapter for more information.) As a result, 2SLS and LIML estimates are not asymptotically efficient. The system methods are 3SLS,

IT3SLS, and FIML. These methods use information concerning the endogenous variables in the system and take into account error covariances across equations and hence are asymptotically efficient in the absence of specification error.

K-class estimation is a class of estimation methods that include the 2SLS, OLS, LIML, and MELO methods as special cases. A  $K$ -value less than 1 is recommended but not required.

MELO is a Bayesian K-class estimator. It yields estimates that can be expressed as a matrix weighted average of the OLS and 2SLS estimates.

The SUR and ITSUR methods use information about contemporaneous correlation among error terms across equations in an attempt to improve the efficiency of parameter estimates.

### **Instrumental Variables and K-Class Estimation Methods**

Instrumental variable methods involve substituting a predicted variable for the endogenous variable  $Y$  when it appears as a regressor. The predicted variables are linear functions of the instrumental variables and the endogenous variable.

The 2SLS method substitutes  $\hat{Y}$  for  $Y$ , which results in consistent estimates. In 2SLS, the instrumental variables are used as regressors to obtain the projected value  $\hat{Y}$ , which is then substituted for  $Y$ . Normally, the predetermined variables of the system are used as the instruments. It is possible to use variables other than predetermined variables from your system of equations as instruments; however, the estimation may not be as efficient. For consistent estimates, the instruments must be uncorrelated with the residual and correlated with the endogenous variable.

K-class estimators are instrumental variable estimators where the first-stage predicted values take a special form:  $Y^* = (1 - k)Y + k\hat{Y}$  for a specified value  $k$ . The probability limit of  $k$  must equal 1 for consistent parameter estimates.

The LIML method results in consistent estimates that are exactly equal to 2SLS estimates when an equation is exactly identified. LIML can be viewed as least-variance ratio estimators or as maximum likelihood estimators. LIML involves minimizing the ratio  $\lambda = (rvar\_eq)/(rvar\_sys)$ , where  $rvar\_eq$  is the residual variance associated with regressing the weighted endogenous variables on all predetermined variables appearing in that equation, and  $rvar\_sys$  is the residual variance associated with regressing weighted endogenous variables on all predetermined variables in the system. The K-class interpretation of LIML is that  $k = \lambda$ . Unlike OLS and 2SLS, where  $k$  is 0 and 1, respectively,  $k$  is stochastic in the LIML method.

The MELO method computes the minimum expected loss estimator. The MELO method computes estimates that "minimize the posterior expectation of generalized quadratic loss functions for structural coefficients of linear structural models" (Judge et al. 1985, 635). Other frequently used K-class estimators may not have finite moments under some commonly encountered circumstances and hence there can be infinite risk relative to quadratic and other loss functions. MELO estimators have finite second moments and hence finite risk.

## Part 2. General Information

One way of comparing K-class estimators is to note that when  $k=1$ , the correlation between regressor and the residual is completely corrected for. In all other cases, it is only partially corrected for.

### **SUR and 3SLS Estimation Methods**

SUR may improve the efficiency of parameter estimates when there is contemporaneous correlation of errors across equations. In practice, the contemporaneous correlation matrix is estimated using OLS residuals. Under two sets of circumstances, SUR parameter estimates are the same as those produced by OLS: when there is no contemporaneous correlation of errors across equations (the estimate of contemporaneous correlation matrix is diagonal); and when the independent variables are the same across equations.

Theoretically, SUR parameter estimates will always be at least as efficient as OLS in large samples, provided that your equations are correctly specified. However, in small samples the need to estimate the covariance matrix from the OLS residuals increases the sampling variability of the SUR estimates, and this effect can cause SUR to be less efficient than OLS. If the sample size is small and the across-equation correlations are small, then OLS should be preferred to SUR. The consequences of specification error are also more serious with SUR than with OLS.

The 3SLS method combines the ideas of the 2SLS and SUR methods. Like 2SLS, the 3SLS method uses  $\hat{Y}$  instead of  $Y$  for endogenous regressors, which results in consistent estimates. Like SUR, the 3SLS method takes the cross-equation error correlations into account to improve large sample efficiency. For 3SLS, the 2SLS residuals are used to estimate the cross-equation error covariance matrix.

The SUR and 3SLS methods can be iterated by recomputing the estimate of the cross-equation covariance matrix from the SUR or 3SLS residuals and then computing new SUR or 3SLS estimates based on this updated covariance matrix estimate. Continuing this iteration until convergence produces ITSUR or IT3SLS estimates.

### **FIML Estimation Method**

The FIML estimator is a system generalization of the LIML estimator. The FIML method involves minimizing the determinant of the covariance matrix associated with residuals of the reduced form of the equation system. From a maximum likelihood standpoint, the LIML method involves assuming that the errors are normally distributed and then maximizing the likelihood function subject to restrictions on a particular equation. FIML is similar, except that the likelihood function is maximized subject to restrictions on all of the parameters in the model, not just those in the equation being estimated. The FIML method is implemented as an instrumental variable method (Hausman 1975).

Note: the RESTRICT, SRESTRICT, TEST, and STEST statements are not supported when the FIML method is used.

### **Choosing a Method for Simultaneous Equations**

A number of factors should be taken into account in choosing an estimation method. Although system methods are asymptotically most efficient in the absence of specification error, system methods are more sensitive to specification error than single equation methods.

In practice, models are never perfectly specified. It is a matter of judgment whether the misspecification is serious enough to warrant avoidance of system methods.

Another factor to consider is sample size. With small samples, 2SLS may be preferred to 3SLS. In general, it is difficult to say much about the small sample properties of K-class estimators because this depends on the regressors used.

LIML and FIML are invariant to the normalization rule imposed but are computationally more expensive than 2SLS or 3SLS.

If the reason for contemporaneous correlation among errors across equations is a common omitted variable, it is not necessarily best to apply SUR. SUR parameter estimates are more sensitive to specification error than OLS. OLS may produce better parameter estimates under these circumstances. SUR estimates are also affected by the sampling variation of the error covariance matrix. There is some evidence from Monte Carlo studies that SUR is less efficient than OLS in small samples.

---

## ANOVA Table for Instrumental Variables Methods

In the instrumental variables methods (2SLS, LIML, K-class, MELO), first-stage predicted values are substituted for the endogenous regressors. As a result, the regression sum of squares (RSS) and the error sum of squares (ESS) do not sum to the total corrected sum of squares for the dependent variable (TSS). The "Analysis of Variance" table printed for the second-stage results serves to display these sums of squares and the mean squares used for the  $F$ -test, but this table is not a variance decomposition in the usual analysis of variance sense.

The  $F$ -test shown in the instrumental variables case is a valid test of the no-regression hypothesis that the true coefficients of all regressors are 0. However, because of the first-stage projection of the regression mean square, this is a Wald-type test statistic, which is asymptotically  $F$  but not exactly  $F$ -distributed in finite samples. Thus, for small samples the  $F$ -test is only approximate when instrumental variables are used.

---

## The $R^2$ Statistics

As explained in the section "ANOVA Table for Instrumental Variables Methods," when instrumental variables are used, the regression sum of squares (RSS) and the error sum of squares (ESS) do not sum to the total corrected sum of squares. In this case, there are several ways that the  $R^2$  statistic can be defined.

The definition of  $R^2$  used by the SYSLIN procedure is

$$R^2 = \frac{\text{RSS}}{\text{RSS} + \text{ESS}}$$

This definition is consistent with the  $F$ -test of the null hypothesis that the true coefficients of all regressors are zero. However, this  $R^2$  may not be a good measure of the goodness of fit of the model.

### **System Weighted $R^2$ and System Weighted Mean Square Error**

The system weighted  $R^2$ , printed for the 3SLS, IT3SLS, SUR, ITSUR, and FIML methods, is computed as follows.

## Part 2. General Information

$$R^2 = Y'WR(X'X)^{-1}R'WY/Y'WY$$

In this equation the matrix  $X'X$  is  $R'WR$ , and  $W$  is the projection matrix of the instruments:

$$W = S^{-1} \otimes Z(Z'Z)^{-1}Z'$$

The matrix  $Z$  is the instrument set,  $R$  is the the regressor set, and  $S$  is the estimated cross-model covariance matrix.

The system weighted MSE, printed for the 3SLS, IT3SLS, SUR, ITSUR, and FIML methods, is computed as follows:

$$MSE = \frac{1}{tdf}(Y'WY - Y'WR(X'X)^{-1}R'WY)$$

In this equation  $tdf$  is the sum of the error degrees of freedom for the equations in the system.

---

## Computational Details

This section discusses various computational details.

### **Computation of Model Crossproduct Matrix**

Model crossproduct matrix  $X'X$  is formed from projected values. For K-class estimation,

$$X = (1 - k)R + kZ(Z'Z)^{-1}Z'R$$

where  $Z$  is the instrument set and  $R$  is the the regressor set. Note that  $k=1$  for the 2SLS method and  $k=0$  for the OLS method.

In the 3SLS, IT3SLS, SUR, and ITSUR methods,  $X'X$  is formed as

$$X'X = R'(S^{-1} \otimes Z(Z'Z)^{-1}Z')R$$

where  $Z$  and  $R$  are as defined previously and  $S$  is an estimate of the cross-equation covariance matrix. For SUR and ITSUR,  $Z$  is the identity matrix.

### **Computation of Standard Errors**

The VARDEF= option in the PROC SYSLIN statement controls the denominator used in calculating the cross-equation covariance estimates and the parameter standard errors and covariances. The values of the VARDEF= option and the resulting denominator are as follows:

- |    |  |
|----|--|
| N  | uses the number of nonmissing observations.  |
| DF | uses the number of nonmissing observations less the degrees of freedom in the model. |



WEIGHT	uses the sum of the observation weights given by the WEIGHTS statement.
WDF	uses the sum of the observation weights given by the WEIGHTS statement less the degrees of freedom in the model.

The VARDEF= option does not affect the model mean square error, root mean square error, or  $R^2$  statistics. These statistics are always based on the error degrees of freedom, regardless of the VARDEF= option. The VARDEF= option also does not affect the dependent variable coefficient of variation (C.V.).

### **Reduced Form Estimates**

The REDUCED option on the PROC SYSLIN statement computes estimates of the reduced form coefficients. The REDUCED option requires that the equation system be square. If there are fewer models than endogenous variables, IDENTITY statements can be used to complete the equation system.

The reduced form coefficients are computed as follows. Represent the equation system, with all endogenous variables moved to the left-hand side of the equations and identities, as

$$\mathbf{B}\mathbf{Y} = \mathbf{\Gamma}\mathbf{X}$$

Here  $\mathbf{B}$  is the estimated coefficient matrix for the endogenous variables  $\mathbf{Y}$ , and  $\mathbf{\Gamma}$  is the estimated coefficient matrix for the exogenous (or predetermined) variables  $\mathbf{X}$ .

The system can be solved for  $\mathbf{Y}$  as follows, provided  $\mathbf{B}$  is square and nonsingular:

$$\mathbf{Y} = \mathbf{B}^{-1}\mathbf{\Gamma}\mathbf{X}$$

The reduced form coefficients are the matrix  $\mathbf{B}^{-1}\mathbf{\Gamma}$ .

### **Uncorrelated Errors Across Equations**

The SDIAG option in the PROC SYSLIN statement computes estimates assuming uncorrelated errors across equations. As a result, when the SDIAG option is used, the 3SLS estimates are identical to 2SLS estimates, and the SUR estimates are the same as the OLS estimates.

### **Over Identification Restrictions**

The OVERID option in the MODEL statement can be used to test for over identifying restrictions on parameters of each equation. The null hypothesis is that the predetermined variables not appearing in any equation have zero coefficients. The alternative hypothesis is that at least one of the assumed zero coefficients is nonzero. The test is approximate and rejects the null hypothesis too frequently for small sample sizes.

The formula for the test is given as follows. Let  $y_i = \beta_i\mathbf{Y}_i + \gamma_i\mathbf{Z}_i + e_i$  be the  $i$ th equation.  $\mathbf{Y}_i$  are the endogenous variables that appear as regressors in the  $i$ th equation, and  $\mathbf{Z}_i$  are the instrumental variables that appear as regressors in the  $i$ th equation. Let  $N_i$  be the number of variables in  $\mathbf{Y}_i$  and  $\mathbf{Z}_i$ .

## Part 2. General Information

Let  $v_i = y_i - \mathbf{Y}_i \hat{\beta}_i$ . Let  $\mathbf{Z}$  represent all instrumental variables,  $T$  be the total number of observations, and  $K$  be the total number of instrumental variables. Define  $\hat{l}$  as follows:

$$\hat{l} = \frac{v_i'(\mathbf{I} - \mathbf{Z}_i(\mathbf{Z}'_i\mathbf{Z}_i)^{-1}\mathbf{Z}'_i)v_i}{v_i'(\mathbf{I} - \mathbf{Z}(\mathbf{Z}'\mathbf{Z})^{-1}\mathbf{Z}')v_i}$$

Then the test statistic

$$\frac{T - K}{K - N_i}(\hat{l} - 1)$$

is distributed approximately as an  $F$  with  $K - N_i$  and  $T - K$  degrees of freedom. Refer to Basman (1960) for more information.

### Fuller's Modification to LIML

The ALPHA= option in the PROC SYSLIN and MODEL statements parameterizes Fuller's modification to LIML. This modification is  $k = \gamma - (\alpha/(n - g))$ , where  $\alpha$  is the value of the ALPHA= option,  $\gamma$  is the LIML  $k$  value,  $n$  is the number of observations, and  $g$  is the number of predetermined variables. Fuller's modification is not used unless the ALPHA= option is specified. Refer to Fuller (1977) for more information.

---

## Missing Values

Observations having a missing value for any variable in the analysis are excluded from the computations.

---

## OUT= Data Set

The output SAS data set produced by the OUT= option in the PROC SYSLIN statement contains all the variables in the input data set and the variables containing predicted values and residuals specified by OUTPUT statements.

The residuals are computed as actual values minus predicted values. Predicted values never use lags of other predicted values, as would be desirable for dynamic simulation. For these applications, PROC SIMLIN is available to predict or simulate values from the estimated equations.

---

## OUTEST= Data Set

The OUTEST= option produces a TYPE=EST output SAS data set containing estimates from the regressions. The variables in the OUTEST= data set are as follows:

BY variables	the BY statement variables are included in the OUTEST= data set
_TYPE_	identifies the estimation type for the observations. The _TYPE_ value INST indicates first-stage regression estimates. Other values indicate the estimation method used: 2SLS indicates two-stage

least squares results, 3SLS indicates three-stage least squares results, LIML indicates limited information maximum likelihood results, and so forth. Observations added by IDENTITY statements have the `_TYPE_` value IDENTITY.

<code>_MODEL_</code>	the model label. The model label is the label specified on the MODEL statement or the dependent variable name if no label is specified. For first-stage regression estimates, <code>_MODEL_</code> has the value FIRST.
<code>_DEPVAR_</code>	the name of the dependent variable for the model
<code>_NAME_</code>	the names of the regressors for the rows of the covariance matrix, if the COVOUT option is specified. <code>_NAME_</code> has a blank value for the parameter estimates observations. The <code>_NAME_</code> variable is not included in the OUTEST= data set unless the COVOUT option is used to output the covariance of parameter estimates matrix.
<code>_SIGMA_</code>	contains the root mean square error for the model, which is an estimate of the standard deviation of the error term. The <code>_SIGMA_</code> variable contains the same values reported as Root MSE in the printed output.
INTERCEPT	the intercept parameter estimates
regressors	the regressor variables from all the MODEL statements are included in the OUTEST= data set. Variables used in IDENTIFY statements are also included in the OUTEST= data set.

The parameter estimates are stored under the names of the regressor variables. The intercept parameters are stored in the variable INTERCEP. The dependent variable of the model is given a coefficient of -1. Variables not in a model have missing values for the OUTEST= observations for that model.

Some estimation methods require computation of preliminary estimates. All estimates computed are output to the OUTEST= data set. For each BY group and each estimation, the OUTEST= data set contains one observation for each MODEL or IDENTITY statement. Results for different estimations are identified by the `_TYPE_` variable.

For example, consider the following statements:

```
proc syslin data=a outest=est 3sls;
  by b;
  endogenous y1 y2;
  instruments x1-x4;
  model y1 = y2 x1 x2;
  model y2 = y1 x3 x4;
  identity x1 = x3 + x4;
run;
```

The 3SLS method requires both a preliminary 2SLS stage and preliminary first stage regressions for the endogenous variable. The OUTEST= data set thus contains 3

different kinds of estimates. The observations for the first-stage regression estimates have the `_TYPE_` value `INST`. The observations for the 2SLS estimates have the `_TYPE_` value `2SLS`. The observations for the final 3SLS estimates have the `_TYPE_` value `3SLS`.

Since there are 2 endogenous variables in this example, there are 2 first-stage regressions and 2 `_TYPE_=INST` observations in the `OUTEST=` data set. Since there are 2 model statements, there are 2 `OUTEST=` observations with `_TYPE_=2SLS` and 2 observations with `_TYPE_=3SLS`. In addition, the `OUTEST=` data set contains an observation with the `_TYPE_` value `IDENTITY` containing the coefficients specified by the `IDENTITY` statement. All these observations are repeated for each `BY`-group in the input data set defined by the values of the `BY` variable `B`.

When the `COVOUT` option is specified, the estimated covariance matrix for the parameter estimates is included in the `OUTEST=` data set. Each observation for parameter estimates is followed by observations containing the rows of the parameter covariance matrix for that model. The row of the covariance matrix is identified by the variable `_NAME_`. For observations that contain parameter estimates, `_NAME_` is blank. For covariance observations, `_NAME_` contains the regressor name for the row of the covariance matrix, and the regressor variables contain the covariances.

See Example 19.1 for an example of the `OUTEST=` data set.

---

## OUTSSCP= Data Set

The `OUTSSCP=` option produces a `TYPE=SSCP` output SAS data set containing sums of squares and cross products. The data set contains all variables used in the `MODEL`, `IDENTITY`, and `VAR` statements. Observations are identified by the variable `_NAME_`.

The `OUTSSCP=` data set can be useful when a large number of observations are to be explored in many different `SYSLIN` runs. The sum-of-squares-and-crossproducts matrix can be saved with the `OUTSSCP=` option and used as the `DATA=` data set on subsequent `SYSLIN` runs. This is much less expensive computationally because `PROC SYSLIN` never reads the original data again. In the step that creates the `OUTSSCP=` data set, include in the `VAR` statement all the variables you expect to use.

---

## Printed Output

The printed output produced by the `SYSLIN` procedure is as follows:

1. If the `SIMPLE` option is used, a table of descriptive statistics is printed showing the sum, mean, sum of squares, variance, and standard deviation for all the variables used in the models.
2. First-stage regression results are printed if the `FIRST` option is specified and an instrumental variables method is used. This shows the regression of each endogenous variable on the variables in the `INSTRUMENTS` list.
3. The results of the second-stage regression are printed for each model. (See "Printed Output for Each Model," which follows.)

4. If a systems method like 3SLS, SUR, or FIML is used, the cross-equation error covariance matrix is printed. This matrix is shown four ways: the covariance matrix itself, the correlation matrix form, the inverse of the correlation matrix, and the inverse of the covariance matrix.
5. If a systems method like 3SLS, SUR, or FIML is used, the system weighted mean square error and system weighted  $R^2$  statistics are printed. The system weighted MSE and  $R^2$  measure the fit of the joint model obtained by stacking all the models together and performing a single regression with the stacked observations weighted by the inverse of the model error variances.
6. If a systems method like 3SLS, SUR, or FIML is used, the final results are printed for each model.
7. If the REDUCED option is used, the reduced form coefficients are printed. This consists of the structural coefficient matrix for the endogenous variables, the structural coefficient matrix for the exogenous variables, the inverse of the endogenous coefficient matrix, and the reduced form coefficient matrix. The reduced form coefficient matrix is the product of the inverse of the endogenous coefficient matrix and the exogenous structural coefficient matrix.

### **Printed Output for Each Model**

The results printed for each model include the "Analysis of Variance" table, the "Parameter Estimates" table, and optional items requested by TEST statements or by options on the MODEL statement.

The printed output produced for each model is described in the following.

The Analysis of Variance table includes the following:

- the model degrees of freedom, sum of squares, and mean square
- the error degrees of freedom, sum of squares, and mean square. The error mean square is computed by dividing the error sum of squares by the error degrees of freedom and is not effected by the VARDEF= option.
- the corrected total degrees of freedom and total sum of squares. Note that for instrumental variables methods the model and error sums of squares do not add to the total sum of squares.
- the  $F$ -ratio, labeled "F Value," and its significance, labeled "PROB>F," for the test of the hypothesis that all the nonintercept parameters are 0
- the root mean square error. This is the square root of the error mean square.
- the dependent variable mean
- the coefficient of variation (C.V.) of the dependent variable
- the  $R^2$  statistic. This  $R^2$  is computed consistently with the calculation of the  $F$  statistic. It is valid for hypothesis tests but may not be a good measure of fit for models estimated by instrumental variables methods.
- the  $R^2$  statistic adjusted for model degrees of freedom, labeled "Adj R-SQ"

The Parameter Estimates table includes the following.

## Part 2. General Information

- estimates of parameters for regressors in the model and the Lagrangian parameter for each restriction specified
- a degrees of freedom column labeled DF. Estimated model parameters have 1 degree of freedom. Restrictions have a DF of -1. Regressors or restrictions dropped from the model due to collinearity have a DF of 0.
- the standard errors of the parameter estimates
- the *t* statistics, which are the parameter estimates divided by the standard errors
- the significance of the *t*-tests for the hypothesis that the true parameter is 0, labeled "Pr > |t|." As previously noted, the significance tests are strictly valid in finite samples only for OLS estimates but are asymptotically valid for the other methods.
- the standardized regression coefficients, if the STB option is specified. This is the parameter estimate multiplied by the ratio of the standard deviation of the regressor to the standard deviation of the dependent variable.
- the labels of the regressor variables or restriction labels

In addition to the Analysis of Variance table and the Parameter Estimates table, the results printed for each model may include the following:

1. If TEST statements are specified, the test results are printed.
2. If the DW option is specified, the Durbin-Watson statistic and first-order autocorrelation coefficient are printed.
3. If the OVERID option is specified, the results of Basmann's test for overidentifying restrictions are printed.
4. If the PLOT option is used, plots of residual against each regressor are printed.
5. If the COVB or CORRB options are specified, the results for each model also include the covariance or correlation matrix of the parameter estimates. For systems methods like 3SLS and FIML, the COVB and CORB output is printed for the whole system after the output for the last model, instead of separately for each model.

The third stage output for 3SLS, SUR, IT3SLS, ITSUR, and FIML does not include the Analysis of Variance table. When a systems method is used, the second stage output does not include the optional output, except for the COVB and CORB matrices.

---

## ODS Table Names

PROC SYSLIN assigns a name to each table it creates. You can use these names to reference the table when using the Output Delivery System (ODS) to select tables and create output data sets. These names are listed in the following table. For more information on ODS, see Chapter 6, "Using the Output Delivery System."

**Table 19.1.** ODS Tables Produced in PROC SYSLIN

<b>ODS Table Name</b>	<b>Description</b>	<b>Option</b>
ANOVA	Summary of the SSE, MSE for the equations	default
AugXPXMat	Model Crossproducts	XPX
AutoCorrStat	Autocorrelation Statistics	default
ConvCrit	Convergence criteria for estimation	default
ConvergenceStatus	Convergence status	default
CorrB	Correlations of parameters	CORRB
CorrResiduals	Correlations of residuals	CORRS
CovB	Covariance of parameters	COVB
CovResiduals	Covariance of residuals	COVS
EndoMat	Endogenous Variables	
Equations	Listing of equations to estimate	default
ExogMat	Exogenous Variables	
FitStatistics	Statistics of Fit	default
InvCorrResiduals	Inverse Correlations of residuals	CORRS
InvCovResiduals	InvCovariance of residuals	COVS
InvEndoMat	Inverse Endogenous Variables	
InvXPX	$X'X$ inverse for System	I
IterHistory	Iteration printing	ITALL/ITPRINT
MissingValues	Missing values generated by the program	default
ModelVars	Name and label for the Model	default
ParameterEstimates	Parameter Estimates	default
RedMat	Reduced Form	REDUCED
SimpleStatistics	Descriptive statistics	SIMPLE
SSCP	Model Crossproducts	
TestResults	Test for Overidentifying Restrictions	
Weight	Weighted Model Statistics	
YPY	Y'Y matrices	USSCP2

## Examples

### Example 19.1. Klein's Model I Estimated with LIML and 3SLS

This example uses PROC SYSLIN to estimate the classic Klein Model I. For a discussion of this model, see Theil (1971). The following statements read the data.

```

*-----Klein's Model I-----*
| By L.R. Klein, Economic Fluctuations in the United States, |
| 1921-1941 (1950), NY: John Wiley. A macro-economic model |
| of the U.S. with three behavioral equations, and several  |
| identities. See Theil, p.456.                             |
*-----*
data klein;
  input year c p w i x wp g t k wsum;
  date=mdy(1,1,year);
  format date monyy.;
  y =c+i+g-t;
  yr =year-1931;
  klag=lag(k);
  plag=lag(p);
  xlag=lag(x);
  label year='Year'
        date='Date'
        c  ='Consumption'
        p  ='Profits'
        w  ='Private Wage Bill'
        i  ='Investment'
        k  ='Capital Stock'
        y  ='National Income'
        x  ='Private Production'
        wsum='Total Wage Bill'
        wp  ='Govt Wage Bill'
        g  ='Govt Demand'
        i  ='Taxes'
        klag='Capital Stock Lagged'
        plag='Profits Lagged'
        xlag='Private Product Lagged'
        yr  ='YEAR-1931';
  datalines;
1920  .  12.7  .  .  44.9  .  .  .  182.8  .
1921  41.9  12.4  25.5  -0.2  45.6  2.7  3.9  7.7  182.6  28.2
1922  45.0  16.9  29.3  1.9  50.1  2.9  3.2  3.9  184.5  32.2
1923  49.2  18.4  34.1  5.2  57.2  2.9  2.8  4.7  189.7  37.0
1924  50.6  19.4  33.9  3.0  57.1  3.1  3.5  3.8  192.7  37.0
1925  52.6  20.1  35.4  5.1  61.0  3.2  3.3  5.5  197.8  38.6
1926  55.1  19.6  37.4  5.6  64.0  3.3  3.3  7.0  203.4  40.7
1927  56.2  19.8  37.9  4.2  64.4  3.6  4.0  6.7  207.6  41.5
1928  57.3  21.1  39.2  3.0  64.5  3.7  4.2  4.2  210.6  42.9
1929  57.8  21.7  41.3  5.1  67.0  4.0  4.1  4.0  215.7  45.3
1930  55.0  15.6  37.9  1.0  61.2  4.2  5.2  7.7  216.7  42.1
1931  50.9  11.4  34.5  -3.4  53.4  4.8  5.9  7.5  213.3  39.3
1932  45.6  7.0  29.0  -6.2  44.3  5.3  4.9  8.3  207.1  34.3
1933  46.5  11.2  28.5  -5.1  45.1  5.6  3.7  5.4  202.0  34.1
1934  48.7  12.3  30.6  -3.0  49.7  6.0  4.0  6.8  199.0  36.6
1935  51.3  14.0  33.2  -1.3  54.4  6.1  4.4  7.2  197.7  39.3
1936  57.7  17.6  36.8  2.1  62.7  7.4  2.9  8.3  199.8  44.2

```



```

1937  58.7  17.3  41.0  2.0  65.0  6.7  4.3  6.7  201.8  47.7
1938  57.5  15.3  38.2 -1.9  60.9  7.7  5.3  7.4  199.9  45.9
1939  61.6  19.0  41.6  1.3  69.5  7.8  6.6  8.9  201.2  49.4
1940  65.0  21.1  45.0  3.3  75.7  8.0  7.4  9.6  204.5  53.0
1941  69.7  23.5  53.3  4.9  88.4  8.5  13.8  11.6  209.4  61.8
;

```

The following statements estimate the Klein model using the limited information maximum likelihood method. In addition, the parameter estimates are written to a SAS data set with the OUTEST= option.

```

proc syslin data=klein outest=b liml;
  endogenous c p w i x wsum k y;
  instruments klag plag xlag wp g t yr;
  consume: model c = p plag wsum;
  invest:  model i = p plag klag;
  labor:   model w = x xlag yr;
run;

proc print data=b; run;

```

The PROC SYSLIN estimates are shown in Output 19.1.1.

**Output 19.1.1.** LIML Estimates

The SYSLIN Procedure						
Limited-Information Maximum Likelihood Estimation						
Model		CONSUME				
Dependent Variable		c				
Label		Consumption				
Analysis of Variance						
Source	DF	Sum of Squares	Mean Square	F Value	Pr > F	
Model	3	854.3541	284.7847	118.42	<.0001	
Error	17	40.88419	2.404952			
Corrected Total	20	941.4295				
Root MSE		1.55079	R-Square	0.95433		
Dependent Mean		53.99524	Adj R-Sq	0.94627		
Coeff Var		2.87209				
Parameter Estimates						
Variable	DF	Parameter Estimate	Standard Error	t Value	Pr >  t	Variable Label
Intercept	1	17.14765	2.045374	8.38	<.0001	Intercept
p	1	-0.22251	0.224230	-0.99	0.3349	Profits
plag	1	0.396027	0.192943	2.05	0.0558	Profits Lagged
wsum	1	0.822559	0.061549	13.36	<.0001	Total Wage Bill

Part 2. General Information

The SYSLIN Procedure						
Limited-Information Maximum Likelihood Estimation						
Model		INVEST				
Dependent Variable		i				
Label		Taxes				
Analysis of Variance						
Source	DF	Sum of Squares	Mean Square	F Value	Pr > F	
Model	3	210.3790	70.12634	34.06	<.0001	
Error	17	34.99649	2.058617			
Corrected Total	20	252.3267				
Root MSE		1.43479	R-Square	0.85738		
Dependent Mean		1.26667	Adj R-Sq	0.83221		
Coeff Var		113.27274				
Parameter Estimates						
Variable	DF	Parameter Estimate	Standard Error	t Value	Pr >  t	Variable Label
Intercept	1	22.59083	9.498146	2.38	0.0294	Intercept
p	1	0.075185	0.224712	0.33	0.7420	Profits
plag	1	0.680386	0.209145	3.25	0.0047	Profits Lagged
klag	1	-0.16826	0.045345	-3.71	0.0017	Capital Stock Lagged

The SYSLIN Procedure						
Limited-Information Maximum Likelihood Estimation						
Model			LABOR			
Dependent Variable			w			
Label			Private Wage Bill			
Analysis of Variance						
Source	DF	Sum of Squares	Mean Square	F Value	Pr > F	
Model	3	696.1485	232.0495	393.62	<.0001	
Error	17	10.02192	0.589525			
Corrected Total	20	794.9095				
Root MSE		0.76781	R-Square	0.98581		
Dependent Mean		36.36190	Adj R-Sq	0.98330		
Coeff Var		2.11156				
Parameter Estimates						
Variable	DF	Parameter Estimate	Standard Error	t Value	Pr >  t	Variable Label
Intercept	1	1.526187	1.320838	1.16	0.2639	Intercept
x	1	0.433941	0.075507	5.75	<.0001	Private Production
xlag	1	0.151321	0.074527	2.03	0.0583	Private Product Lagged
yr	1	0.131593	0.035995	3.66	0.0020	YEAR-1931

The OUTEST= data set is shown in part in Output 19.1.2. Note that the data set contains the parameter estimates and root mean square errors, `_SIGMA_`, for the first stage instrumental regressions as well as the parameter estimates and  $\sigma$  for the LIML estimates for the three structural equations.

#### Output 19.1.2. The OUTEST= Data Set

Obs	_TYPE_	_STATUS_	_MODEL_	_DEPVAR_	_SIGMA_	Intercept	klag	plag					
1	LIML	0 Converged	CONSUME	c	1.55079	17.1477	.	0.39603					
2	LIML	0 Converged	INVEST	i	1.43479	22.5908	-0.16826	0.68039					
3	LIML	0 Converged	LABOR	w	0.76781	1.5262	.	.					
Obs	xlag	wp	g	t	yr	c	p	w	i	x	wsum	k	y
1	.	.	.	.	.	-1	-0.22251	.	.	.	0.82256	.	.
2	.	.	.	.	.	.	0.07518	.	-1	.	.	.	.
3	0.15132	.	.	.	0.13159	.	.	-1	.	0.43394	.	.	.

The following statements estimate the model using the 3SLS method. The reduced form estimates are produced by the `REDUCED` option; `IDENTITY` statements are used to make the model complete.

```
proc syslin data=klein 3sls reduced;
  endogenous c p w i x wsum k y;
  instruments klag plag xlag wp g t yr;
  consume: model    c = p plag wsum;
```

Part 2. General Information

```

invest:  model    i = p plag klag;
labor:   model    w = x xlag yr;
product: identity x = c + i + g;
income:  identity y = c + i + g - t;
profit:  identity p = y - w;
stock:   identity k = klag + i;
wage:    identity wsum = w + wp;
run;

```

The preliminary 2SLS results and estimated cross-model covariance matrix are not shown. The 3SLS estimates are shown in Output 19.1.3. The reduced form estimates are shown in Output 19.1.4.

**Output 19.1.3. 3SLS Estimates**

The SYSLIN Procedure						
Three-Stage Least Squares Estimation						
System Weighted MSE		5.9342				
Degrees of freedom		51				
System Weighted R-Square		0.9550				
Model		CONSUME				
Dependent Variable		c				
Label		Consumption				
Parameter Estimates						
Variable	DF	Parameter Estimate	Standard Error	t Value	Pr >  t	Variable Label
Intercept	1	16.44079	1.449925	11.34	<.0001	Intercept
p	1	0.124890	0.120179	1.04	0.3133	Profits
plag	1	0.163144	0.111631	1.46	0.1621	Profits Lagged
wsum	1	0.790081	0.042166	18.74	<.0001	Total Wage Bill

The SYSLIN Procedure						
Three-Stage Least Squares Estimation						
Model		INVEST				
Dependent Variable		i				
Label		Taxes				
Parameter Estimates						
Variable	DF	Parameter Estimate	Standard Error	t Value	Pr >  t	Variable Label
Intercept	1	28.17785	7.550853	3.73	0.0017	Intercept
p	1	-0.01308	0.179938	-0.07	0.9429	Profits
plag	1	0.755724	0.169976	4.45	0.0004	Profits Lagged
klag	1	-0.19485	0.036156	-5.39	<.0001	Capital Stock Lagged

```

The SYSLIN Procedure
Three-Stage Least Squares Estimation

Model                                LABOR
Dependent Variable                    w
Label                                Private Wage Bill

Parameter Estimates

Variable      Parameter Standard      Variable
              DF Estimate   Error t Value Pr > |t| Label
Intercept    1  1.797218 1.240203    1.45  0.1655 Intercept
x            1  0.400492 0.035359   11.33 <.0001 Private Production
xlag        1  0.181291 0.037965    4.78  0.0002 Private Product Lagged
yr          1  0.149674 0.031048    4.82  0.0002 YEAR-1931
    
```

**Output 19.1.4. Reduced Form Estimates**

```

The SYSLIN Procedure
Three-Stage Least Squares Estimation

Endogenous Variables

              c              p              w              i
CONSUME      1      -0.12489              0              0
INVEST       0      0.013079              0              1
LABOR        0              0              1              0
PRODUCT     -1              0              0             -1
INCOME      -1              0              0             -1
PROFIT       0              1              1              0
STOCK        0              0              0             -1
WAGE         0              0             -1              0

Endogenous Variables

              x              wsum              k              y
CONSUME      0      -0.79008              0              0
INVEST       0              0              0              0
LABOR      -0.40049              0              0              0
PRODUCT      1              0              0              0
INCOME       0              0              0              1
PROFIT       0              0              0             -1
STOCK        0              0              1              0
WAGE         0              1              0              0
    
```

Part 2. General Information

The SYSLIN Procedure  
Three-Stage Least Squares Estimation

Exogenous Variables

	Intercept	plag	klag	xlag
CONSUME	16.44079	0.163144	0	0
INVEST	28.17785	0.755724	-0.19485	0
LABOR	1.797218	0	0	0.181291
PRODUCT	0	0	0	0
INCOME	0	0	0	0
PROFIT	0	0	0	0
STOCK	0	0	1	0
WAGE	0	0	0	0

Exogenous Variables

	yr	g	t	wp
CONSUME	0	0	0	0
INVEST	0	0	0	0
LABOR	0.149674	0	0	0
PRODUCT	0	1	0	0
INCOME	0	1	-1	0
PROFIT	0	0	0	0
STOCK	0	0	0	0
WAGE	0	0	0	1

The SYSLIN Procedure  
Three-Stage Least Squares Estimation

Inverse Endogenous Variables

	CONSUME	INVEST	LABOR	PRODUCT
c	1.634654	0.634654	1.095657	0.438802
p	0.972364	0.972364	-0.34048	-0.13636
w	0.649572	0.649572	1.440585	0.576943
i	-0.01272	0.987282	0.004453	0.001783
x	1.621936	1.621936	1.10011	1.440585
wsum	0.649572	0.649572	1.440585	0.576943
k	-0.01272	0.987282	0.004453	0.001783
y	1.621936	1.621936	1.10011	0.440585

Inverse Endogenous Variables

	INCOME	PROFIT	STOCK	WAGE
c	0.195852	0.195852	0	1.291509
p	1.108721	1.108721	0	0.768246
w	0.072629	0.072629	0	0.513215
i	-0.0145	-0.0145	0	-0.01005
x	0.181351	0.181351	0	1.281461
wsum	0.072629	0.072629	0	1.513215
k	-0.0145	-0.0145	1	-0.01005
y	1.181351	0.181351	0	1.281461

The SYSLIN Procedure				
Three-Stage Least Squares Estimation				
Reduced Form				
	Intercept	plag	klag	xlag
c	46.7273	0.746307	-0.12366	0.198633
p	42.77363	0.893474	-0.18946	-0.06173
w	31.57207	0.596871	-0.12657	0.261165
i	27.6184	0.744038	-0.19237	0.000807
x	74.3457	1.490345	-0.31603	0.19944
wsum	31.57207	0.596871	-0.12657	0.261165
k	27.6184	0.744038	0.80763	0.000807
y	74.3457	1.490345	-0.31603	0.19944

Reduced Form				
	yr	g	t	wp
c	0.163991	0.634654	-0.19585	1.291509
p	-0.05096	0.972364	-1.10872	0.768246
w	0.215618	0.649572	-0.07263	0.513215
i	0.000667	-0.01272	0.014501	-0.01005
x	0.164658	1.621936	-0.18135	1.281461
wsum	0.215618	0.649572	-0.07263	1.513215
k	0.000667	-0.01272	0.014501	-0.01005
y	0.164658	1.621936	-1.18135	1.281461

## Example 19.2. Grunfeld's Model Estimated with SUR

The following example was used by Zellner in his classic 1962 paper on seemingly unrelated regressions. Different stock prices often move in the same direction at a given point in time. The SUR technique may provide more efficient estimates than OLS in this situation.

The following statements read the data. (The prefix GE stands for General Electric and WH stands for Westinghouse.)

```
*-----Zellner's Seemingly Unrelated Technique-----*
| A. Zellner, "An Efficient Method of Estimating Seemingly
| Unrelated Regressions and Tests for Aggregation Bias,"
| JASA 57(1962) pp.348-364
|
| J.C.G. Boot, "Investment Demand: an Empirical Contribution
| to the Aggregation Problem," IER 1(1960) pp.3-30.
|
| Y. Grunfeld, "The Determinants of Corporate Investment,"
| Unpublished thesis, Chicago, 1958
|-----*
data grunfeld;
  input year ge_i ge_f ge_c wh_i wh_f wh_c;
  label ge_i = 'Gross Investment, GE'
        ge_c = 'Capital Stock Lagged, GE'
        ge_f = 'Value of Outstanding Shares Lagged, GE'
        wh_i = 'Gross Investment, WH'
        wh_c = 'Capital Stock Lagged, WH'
        wh_f = 'Value of Outstanding Shares Lagged, WH';
```

## Part 2. General Information

```
      datalines;
1935    33.1    1170.6    97.8    12.93    191.5    1.8
1936    45.0    2015.8    104.4    25.90    516.0    .8
1937    77.2    2803.3    118.0    35.05    729.0    7.4
1938    44.6    2039.7    156.2    22.89    560.4    18.1
1939    48.1    2256.2    172.6    18.84    519.9    23.5
1940    74.4    2132.2    186.6    28.57    628.5    26.5
1941   113.0    1834.1    220.9    48.51    537.1    36.2
1942    91.9    1588.0    287.8    43.34    561.2    60.8
1943    61.3    1749.4    319.9    37.02    617.2    84.4
1944    56.8    1687.2    321.3    37.81    626.7    91.2
1945    93.6    2007.7    319.6    39.27    737.2    92.4
1946   159.9    2208.3    346.0    53.46    760.5    86.0
1947   147.2    1656.7    456.4    55.56    581.4   111.1
1948   146.3    1604.4    543.4    49.56    662.3   130.6
1949    98.3    1431.8    618.3    32.04    583.8   141.8
1950    93.5    1610.5    647.4    32.24    635.2   136.7
1951   135.2    1819.4    671.3    54.38    723.8   129.7
1952   157.3    2079.7    726.1    71.78    864.1   145.5
1953   179.5    2371.6    800.3    90.08   1193.5   174.8
1954   189.6    2759.9    888.9    68.60   1188.9   213.5
;
```

The following statements compute the SUR estimates for the Grunfeld model.

```
proc syslin data=grunfeld sur;
  ge:      model ge_i = ge_f ge_c;
  westing: model wh_i = wh_f wh_c;
run;
```

The PROC SYSLIN output is shown in Output 19.2.1.



## Output 19.2.1. PROC SYSLIN Output for SUR

The SYSLIN Procedure						
Ordinary Least Squares Estimation						
Model		GE				
Dependent Variable		ge_i				
Label		Gross Investment, GE				
Analysis of Variance						
Source	DF	Sum of Squares	Mean Square	F Value	Pr > F	
Model	2	31632.03	15816.02	20.34	<.0001	
Error	17	13216.59	777.4463			
Corrected Total	19	44848.62				
Root MSE		27.88272	R-Square	0.70531		
Dependent Mean		102.29000	Adj R-Sq	0.67064		
Coeff Var		27.25850				
Parameter Estimates						
Variable	DF	Parameter Estimate	Standard Error	t Value	Pr >  t	Variable Label
Intercept	1	-9.95631	31.37425	-0.32	0.7548	Intercept
ge_f	1	0.026551	0.015566	1.71	0.1063	Value of Outstanding Shares Lagged, GE
ge_c	1	0.151694	0.025704	5.90	<.0001	Capital Stock Lagged, GE

Part 2. General Information

The SYSLIN Procedure						
Ordinary Least Squares Estimation						
Model			WESTING			
Dependent Variable			wh_i			
Label			Gross Investment, WH			
Analysis of Variance						
Source	DF	Sum of Squares	Mean Square	F Value	Pr > F	
Model	2	5165.553	2582.776	24.76	<.0001	
Error	17	1773.234	104.3079			
Corrected Total	19	6938.787				
Root MSE		10.21312	R-Square	0.74445		
Dependent Mean		42.89150	Adj R-Sq	0.71438		
Coeff Var		23.81153				
Parameter Estimates						
Variable	DF	Parameter Estimate	Standard Error	t Value	Pr >  t	Variable Label
Intercept	1	-0.50939	8.015289	-0.06	0.9501	Intercept
wh_f	1	0.052894	0.015707	3.37	0.0037	Value of Outstanding Shares Lagged, WH
wh_c	1	0.092406	0.056099	1.65	0.1179	Capital Stock Lagged, WH

```

The SYSLIN Procedure
Seemingly Unrelated Regression Estimation

Cross Model Covariance

      GE          WESTING
GE      777.446    207.587
WESTING 207.587    104.308

Cross Model Correlation

      GE          WESTING
GE      1.00000    0.72896
WESTING 0.72896    1.00000

Cross Model Inverse Correlation

      GE          WESTING
GE      2.13397    -1.55559
WESTING -1.55559    2.13397

Cross Model Inverse Covariance

      GE          WESTING
GE      0.002745   -.005463
WESTING -.005463   0.020458
    
```

```

The SYSLIN Procedure
Seemingly Unrelated Regression Estimation

System Weighted MSE          0.9719
Degrees of freedom           34
System Weighted R-Square     0.6284

Model                          GE
Dependent Variable             ge_i
Label                          Gross Investment, GE

Parameter Estimates

Variable      Parameter Standard      Variable
              DF Estimate   Error t Value Pr > |t| Label
Intercept    1  -27.7193 29.32122  -0.95  0.3577 Intercept
ge_f         1   0.038310 0.014415   2.66  0.0166 Value of Outstanding Shares
              Lagged, GE
ge_c         1   0.139036 0.024986   5.56  <.0001 Capital Stock Lagged, GE
    
```

The SYSLIN Procedure						
Seemingly Unrelated Regression Estimation						
Model			WESTING			
Dependent Variable			wh_i			
Label			Gross Investment, WH			
Parameter Estimates						
Variable	DF	Parameter Estimate	Standard Error	t Value	Pr >  t	Variable Label
Intercept	1	-1.25199	7.545217	-0.17	0.8702	Intercept
wh_f	1	0.057630	0.014546	3.96	0.0010	Value of Outstanding Shares Lagged, WH
wh_c	1	0.063978	0.053041	1.21	0.2443	Capital Stock Lagged, WH

## References

- Basman, R.L. (1960), "On Finite Sample Distributions of Generalized Classical Linear Identifiability Test Statistics," *Journal of the American Statistical Association*, 55, 650-659.
- Fuller, W.A. (1977), "Some Properties of a Modification of the Limited Information Estimator," *Econometrica*, 45, 939-952.
- Hausman, J.A. (1975), "An Instrumental Variable Approach to Full Information Estimators for Linear and Certain Nonlinear Econometric Models," *Econometrica*, 43, 727-738.
- Johnston, J. (1984), *Econometric Methods*, Third Edition, New York: McGraw-Hill Book Company.
- Judge, George G., W. E. Griffiths, R. Carter Hill, Helmut Lutkepohl, and Tsoung-Chao Lee (1985), *The Theory and Practice of Econometrics*, Second Edition, New York: John Wiley & Sons, Inc.
- Maddala, G.S. (1977), *Econometrics*, New York: McGraw-Hill Book Company.
- Park, S.B. (1982), "Some Sampling Properties of Minimum Expected Loss (MELO) Estimators of Structural Coefficients," *Journal of the Econometrics*, 18, 295-311.
- Pindyck, R.S. and Rubinfeld, D.L. (1981), *Econometric Models and Economic Forecasts*, Second Edition, New York: McGraw-Hill Book Company.
- Pringle, R.M. and Raynor, A.A. (1971), *Generalized Inverse Matrices with Applications to Statistics*, New York: Hafner Publishing Company.
- Rao, P. (1974), "Specification Bias in Seemingly Unrelated Regressions," in *Essays in Honor of Tinbergen*, Volume 2, New York: International Arts and Sciences Press.
- Savin, N.E. and White, K.J. (1978), "Testing for Autocorrelation with Missing Observations," *Econometrics*, 46, 59-66.

- Theil, H. (1971), *Principles of Econometrics*, New York: John Wiley & Sons, Inc.
- Zellner, A. (1962), "An Efficient Method of Estimating Seemingly Unrelated Regressions and Tests for Aggregation Bias," *Journal of the American Statistical Association*, 57, 348-368.
- Zellner, A. (1978), "Estimation of Functions of Population Means and Regression Coefficients: A Minimum Expected Loss (MELO) Approach," *Journal of the Econometrics*, 8, 127-158.
- Zellner, A. and Park, S. (1979), "Minimum Expected Loss (MELO) Estimators for Functions of Parameters and Structural Coefficients of Econometric Models," *Journal of the American Statistical Association*, 74, 185-193.

The correct bibliographic citation for this manual is as follows: SAS Institute Inc., *SAS/ETS User's Guide, Version 8*, Cary, NC: SAS Institute Inc., 1999. 1546 pp.

**SAS/ETS User's Guide, Version 8**

Copyright © 1999 by SAS Institute Inc., Cary, NC, USA.

ISBN 1-58025-489-6

All rights reserved. Printed in the United States of America. No part of this publication may be reproduced, stored in a retrieval system, or transmitted, in any form or by any means, electronic, mechanical, photocopying, or otherwise, without the prior written permission of the publisher, SAS Institute Inc.

**U.S. Government Restricted Rights Notice.** Use, duplication, or disclosure of the software by the government is subject to restrictions as set forth in FAR 52.227-19 Commercial Computer Software-Restricted Rights (June 1987).

SAS Institute Inc., SAS Campus Drive, Cary, North Carolina 27513.

1st printing, October 1999

SAS® and all other SAS Institute Inc. product or service names are registered trademarks or trademarks of SAS Institute Inc. in the USA and other countries.® indicates USA registration.

Other brand and product names are registered trademarks or trademarks of their respective companies.

The Institute is a private company devoted to the support and further development of its software and related services.